DIABETES, GLYCOSURIA, AND PROTEINURIA IN A JAPANESE COHORT FOLLOWED FOR 20 YEARS

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JOEL B. BRODSKY, Ph.D.
DIRK F. MOORE, Ph.D.
RYOSO KAWATE, M.D. 川手亮三
HOWARD B. HAMILTON, M.D.

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RERF Departments of Epidemiology & Statistics¹ and Clinical Laboratories,³ and 2nd Department of Internal Medicine, Hiroshima University School of Medicine² 放影研疫学統計部,¹ 臨床検査部,³ 広島大学医学部第 2 内科学教室²

SUMMARY

An analysis of noninsulin-dependent diabetes mellitus, glycosuria, and proteinuria in a Japanese cohort, which has been followed for 20 years, is presented. For the last 14 years of follow-up in Hiroshima, the sex-specific incidence rates of diabetes were approximately constant at 1.2-1.4 cases/100 per 2-year interval for males (p>.50) and at 0.5-0.6 cases/100 per 2-year interval for females (p>.20). The diagnosis of diabetes depended on first screening for glycosuria and then administering an oral glucose tolerance test only to those with glycosuria. The relatively greater prevalence of glycosuria and diabetes among males has also been noted by others using similar methods. It has been reported elsewhere that if the glucose tolerance test were used to screen for diabetes, the sex difference may not be apparent. The prevalence of diabetes increased over the course of study (p<.001), and the increase was attributed to aging of the cohort. The size of the undiagnosed diabetic population was estimated and the observed prevalence rate was determined to be about 75% of the actual prevalence rate at the end of 8 years of follow-up and to be about 89% of the actual prevalence rate at the end of 20 years of follow-up. Proteinuria prevalence was 4.3% in nondiabetics, 23.5% in Hiroshima diabetics, and 18.7% in Nagasaki diabetics. The differences in prevalence rates between nondiabetics and diabetics were highly significant (p<.001).

要約

日本人コホートにおけるインシュリン非依存型糖尿病, 糖尿及び蛋白尿に関する20年間の経過観察の解析結 果を報告する. 広島における過去14年間の経過観察 では,2年ごとの性別糖尿病発生率はほぼ一定で, 男性は100例当たり1.2-1.4例(p>.50), 女性は 100例当たり0.5-0.6例(p>.20)であった. 糖尿病 の診断に当たっては、まず糖尿のスクリーニングを 行った後, 糖尿が認められる者に対してのみ経口 ブドウ糖負荷試験を行い, その結果に基づいて診断 を下した. 糖尿及び糖尿病が男性に好発することは, 同様の方法を用いたほかの研究者によっても認められ ている. 糖尿病のスクリーニングにブドウ糖負荷試験 を用いるならば, 糖尿病有病率は男女間で明確な差 が得られないであろうとほかの研究で報告されている. 調査の実施期間中,糖尿病の有病率は増加したが (p<.001), これはコホートの老齢化に帰因すると 考えられた. 診断の下されていない糖尿病症例集団 の大きさを推定し、経過観察開始の8年後では、 観察有病率は実際の有病率の約75%,20年後では 約89%であると判明した.蛋白尿有病率は非糖尿病 患者では4.3%, 広島の糖尿病患者では23.5%, 長崎の糖尿病患者では18.7%であった. 非糖尿病 患者と糖尿病患者における蛋白尿有病率の差は極め て有意であった(p<.001).

^{*}Department of Biostatistics, School of Public Health and Community Medicine, University of Washington; Former Visiting Research Fellow

Washington 大学公衆衛生·地域医療学部生物統計部門, 元来所研究員

^{**}Kawate Internal Medicine Clinic; Former Visiting Research Associate 川手内科クリニック, 元来所研究員

INTRODUCTION

The etiologic factors associated with diabetes mellitus have fascinated medical practitioners since the early days of human history. Perhaps as much as any disease known to man, the morbidity rates of diabetes mellitus and its complications differ by population. There are numerous examples of vastly differing population prevalence rates and contradictory male to female sex ratios.¹

Surprisingly, even though diabetes is intensely studied in Japan, there have been few long-term followup studies of diabetes prevalence in large Japanese populations. The present report gives the results of such a study.

The subset of the Japanese atomic bomb survivors, who comprise the Adult Health Study (AHS) cohort, provides an excellent population in which to study the epidemiology of noninsulin-dependent diabetes mellitus (NIDDM). The size of the original cohort was approximately 20,000 and the members of the cohort were offered a general medical examination every two years. Therefore, it is possible to investigate the incidence and prevalence rates of diabetes in this group. Rudnick and Anderson,² and Freedman et al³ studied this cohort and described a predominance of male diabetics and large differences between Hiroshima and Nagasaki in diabetes morbidity rates. The last report which gave morbidity rates for diabetes in the AHS was that of Belsky et al.4 Reported here are the data on diabetes for the period 1958-78.

MATERIALS AND METHODS

The Population. The AHS is a cohort study designed to investigate the long-term health status of a subgroup of the Japanese A-bomb survivors and a suitably chosen control group. The approximately 20,000 individuals in the AHS cohort were selected according to information supplied in the 1950 Japanese National Census Supplementary Schedules and for each A-bomb survivor exposed in Hiroshima or Nagasaki at the time of the bomb. The cohort was originally comprised of four groups of equal size: inner proximally exposed (less than 2,000 m from the hypocenter) with major radiation symptoms, inner proximally exposed without symptoms, distally exposed (3,000-3,999 m), and nonexposed controls. The details of the selection process and an analysis of some sociological characteristics of the cohort have been given elsewhere.5

緒言

糖尿病に関連する病因は、人類の歴史が始まって以来、医師の注目を集めてきた. 恐らく、既知のあらゆる病気と同様、糖尿病とその合併症の罹病率は集団によって大きく異なる. 集団有病率が非常に異なった例、また男女比が逆転する例も多い.1

驚くべきことに日本では、糖尿病が集中的に研究されている場合でも、大規模な日本人集団の糖尿病 有病率について長期的な追跡調査がほとんど行われていない.本報では、そのような調査結果について述べる.

成人健康調査(AHS)コホートを構成する日本人原爆被爆者の小集団は、インシュリン非依存型糖尿病の疫学研究を行うには非常に有用な集団である。当初の集団は約20,000人であり、対象者は2年ごとに一般的な医学的検診を受けている。したがって、この集団における糖尿病の発生率及び有病率を調査することは可能である。Rudnick及びAnderson、2並びにFreedmanら3はこの集団について調査し、糖尿病は男性に多く、また糖尿病罹病率は広島・長崎間で大きな差のあることを報告した。AHS集団における糖尿病の罹病率について最後の報告を行ったのはBelskyら4である。今回報告するのは、1958-78年の糖尿病のデータである。

材料及び方法

調査集団、AHSは、日本人原爆被爆者の小集団及びその対照群の健康状態を長期的に調べる目的で計画されたコホート研究である。AHSコホートの約20,000人の対象者は、1950年の国勢調査付帯調査の資料、並びに広島及び長崎の各原爆被爆者に関する資料に基づいて抽出したものである。このコホートは、最初同規模な4群で構成された。すなわち、主要な放射線症状を示す近距離被爆群(爆心地から2,000m以内)、放射線症状を示さない近距離被爆群、遠距離被爆群(3,000-3,999m)及び非被爆の対照群である。抽出過程の詳細とコホートの社会学的特徴の解析については、別に報告されている。5

The AHS subjects have been offered general medical examinations on a biennial basis since the study began in 1958. The cohort is thought to be fairly representative of the general populations of Hiroshima and Nagasaki in 1950, except for the possible effects due to radiation exposure. Previous analyses of diabetes²⁻⁴ and recent mortality studies^{6,7} have failed to reveal any relationship between diabetes and radiation exposure. Thus, if there are no effects of radiation on diabetes then it is possible to estimate 2-year diabetes incidence and prevalence rates in this cohort and to examine how these rates change in time.

Diabetes Prescreening Procedure. A routine urinalysis, including tests for glycosuria and proteinuria, comprises part of the standard biennial examination. If trace or greater glycosuria was found in a subject who had not been previously diagnosed as diabetic, the subject usually was advised to return for a blood glucose test. In this manner, routine urinalysis served as the principal diabetes prescreening procedure. A drawback of this procedure is that diabetics are not always glycosuric. Hence, in each examination cycle undiagnosed diabetics who were not glycosuric at the time of examination generally did not receive a blood glucose test. However, any individual who requested a blood glucose test, whether he or she had been diagnosed as diabetic or not, was offered such a test.

Measurements. Height was measured without shoes and weight was measured with the subject wearing only a dressing gown, and the weight of the gown was subtracted. Ideal body weight (IBW) was expressed as a percentage based upon the Japanese height-weight tables of Minowa et al.8 Obesity was defined as a body weight in excess of 115% IBW and nonobesity was defined as a body weight less than 115% IBW. Classification of an individual with an IBW of exactly 115% was determined from the trend (upward or downward) observed in IBW readings from previous and subsequent examinations. Family histories of diabetes, including information on relatives beyond first degree, were obtained from questionnaires completed during cycle 3 (1962-64) and cycle 4 (1964-66) examinations. An indicated family history on either questionnaire was deemed sufficient for classifying a family history as positive.

AHS 対象者は、1958年に調査が開始されて以来、2年ごとに一般的な医学的検診を受けている。この集団は、放射線被曝による影響を除けば、1950年における広島及び長崎の一般集団をよく代表するものと考えられる。糖尿病に関する過去の解析²⁻⁴ 及び最近の死亡率調査^{6,7}では、糖尿病と放射線被曝との間に何ら関係は認められなかった。したがって、糖尿病に放射線の影響が認められない場合、この集団における2年間の糖尿病発生率及び有病率を推定し、これらの率がどのような経時的変化を示すかを調べることが可能になる。

糖尿病の前スクリーニング手順.2年ごとの標準的な検診の一部として、糖尿及び蛋白尿検査を含む通常の尿検査を行う.以前に糖尿病と診断されていない対象者に微量以上の糖尿が認められた場合、普通、その対象者に血糖検査のために再来所することを勧めた.このように、通常の尿検査が糖尿病前スクリーニングの主な手順であった.この手順の欠点は、糖尿病患者が常に糖尿を伴うとは限らないことにある.したがって、各診察周期の検診時に糖尿が認められず、糖尿病と診断されなかった対象者の大部分は血糖検査を受けていない.しかし血糖検査を希望する対象者には、糖尿病の診断の有無にかかわらず、この検査を行った.

測定. 靴を履かない状態で身長を測定し、ガウンを着けただけで体重を測定し、ガウンの重さを差し引いた. 理想体重 (IBW) は箕輪ら8の日本人の身長・体重表に基づきパーセントで表示した. 肥満は、体重が115% IBW を超えるもの、また非肥満は115% IBW 未満のものと定義した. IBW 115%丁度の対象者の分類は、過去及びその後の検査の IBW 記録にみられる傾向(増加中あるいは減少中)に基づいて決定した. 1親等以上の親族についての情報を含む糖尿病の家族歴は、第3(1962-64年)及び第4(1964-66年)診察周期の間に完了した質問票から得た. 両周期の質問票に示された家族歴は、いずれも家族歴を陽性と分類するのに十分であるとみなした.

Laboratory Measurements. Urinalyses were performed without consideration of the length of time from a subject's previous meal, a period which was assumed to be 2-3 hours in duration. Until June 1960, Benedict's reagent was used for determining glycosuria, and since July 1960 urine was tested with Combistix. The urine was initially tested for protein with sulfosalicylic acid and positive results confirmed by the nitric acid ring test. For cycle 1 (1958-60) examination, nitric acid ring test results were recorded in Hiroshima, but sulfosalicylic acid test results were recorded in Nagasaki. Beginning in July 1960, Combistix was used to test for protein in both cities, and positive results were confirmed with nitric acid. Combistix results were recorded in Nagasaki and nitric acid ring test results were recorded in Hiroshima. Beginning on 23 October 1970 in Hiroshima and 12 November 1970 in Nagasaki, a commercial preparation called Trelan G was used for all oral glucose tolerance tests (OGTT). Prior to these dates, on the morning of each administered OGTT a loading solution was prepared in the laboratory.9 All blood glucose tests were reviewed by the authors and all known postprandial tests were removed from consideration for this study.

Diagnosis of Diabetes. A diagnosis of diabetes was made using the criteria recommended by the National Diabetes Data Group (NDDG), 10 including the requirement of two positive blood glucose tests (the 'first time' positive test and the 'confirmatory' positive test, respectively). A first-time blood glucose test was considered positive if a fasting blood sugar (FBS) test was at least 140 mg/100 ml or if the 1-hour value of a 50 g OGTT was in excess of 200 mg/100 ml and the 2-hour value was in excess of 180 mg/100 ml. For most OGTT the fasting, 1-hour, and 2-hour values were recorded. medical records of each individual who died within one year of a first-time positive blood glucose test were examined for disorders which might have been responsible for the observed glucose intolerance. If such conditions existed, the individual was not deemed diabetic. Also, any individual with a known condition which might have caused a glucose intolerance was not considered diabetic. There were no known instances of either gestational diabetes or insulin-dependent diabetes mellitus (IDDM).

An attempt was made to confirm all first-time positive subjects with a slightly less stringent criterion 上記ほど厳しくない基準により、初回の検査で

臨床検査. 対象者が最後にとった食事からの時間を 考慮に入れずに尿検査を実施した. これは2~3時間 と思われる. 1960年6月まではベネディクト試薬を用い て糖尿の測定を行い、1960年7月以降は、Combistix を用いて尿検査を行った. 尿検査では、まずスルホ サリチル酸で蛋白質の検査を行い, 陽性のものは 硝酸リングテストにより確認した. 第1診察周期 (1958-60年) については、広島では硝酸リングテスト の結果を記録したが、長崎ではスルホサリチル酸検査 の結果を記録した。1960年7月から Combistix を用い て両市の蛋白質検査を行い、結果が陽性のものは 硝酸により確認した. 長崎では Combistix の結果を 記録し,広島では硝酸リングテストの結果を記録した. 広島では1970年10月23日,長崎では1970年11月12日 から, すべての経口ブドウ糖負荷試験(OGTT)に Trelan Gと呼ばれる市販の製剤を使用した. それ 以前には、OGTT を実施する朝ごとに研究室で負荷 溶液を調製していた.9 著者らは, 血糖検査をすべて 検討し, 食後に行ったことがわかった検査はすべて この研究の対象外とした.

糖尿病の診断. 糖尿病は, 2種類の陽性検査(それ ぞれ初回の陽性検査及び確認陽性検査)の必要条件 を含む米国 National Diabetes Data Group (NDDG)10 の勧告基準に基づいて診断した. 初回の血糖検査で, 空腹時血糖 (FBS) の値が少なくとも 140 mg/100 ml を示すか, 又は50g OGTT の1時間の値が200 mg/ 100 ml を超え, 2時間の値が180 mg/100 ml を超える 場合は、陽性とみなした. ほとんどの OGTT について 空腹時, 1時間及び2時間の値を記録した. 初回 陽性血糖検査から1年以内に死亡した各対象者に ついて、観察されたブドウ糖不耐症の原因と考えられ る疾患があるか医学記録を調べた. このような状態が 認められた場合は, その対象者を糖尿病とはみなさ なかった. また, ブドウ糖不耐症の原因となったと 考えられる疾患がわかっている対象者も糖尿病とみな さなかった. 妊娠糖尿病及びインシュリン依存型 糖尿病(IDDM)はいずれも症例がなかった.

than that given above. A fasting blood glucose of at least 135 mg/100 ml or both the 1-hour and 2hour OGTT values in excess of 180 mg/100 ml was considered sufficient for confirmation of a diagnosis of diabetes. Also, 17 cases in Hiroshima and 4 cases in Nagasaki were confirmed using test results obtained prior to the diagnosis cycle. Both of these criteria are more strict than those used in Japan 11 during the course of this study and are more strict than those used in prior studies2-4 of this cohort. The medical records of subjects without confirmatory test results were examined for clinical confirmation or contradiction of a diagnosis. The individuals with neither confirmatory nor contradictory evidence were taken as diabetic. In Nagasaki, 104 cases were reviewed, of which 85 had a clinical history of diabetes and 19 had contradictory evidence. In Hiroshima, 244 cases were reviewed, of which 146* were estimated to have a clinical history of diabetes, 84* were estimated to have contradictory evidence, and 14* were estimated to have neither.

In summary, the "confirmed diabetics" were all patients with at least one of the following:

- a) at least two positive blood glucose tests, without a known (nondiabetic) condition which might account for the positive tests; or
- b) one positive blood glucose test and a clinical history of diabetes; or
- c) one positive blood glucose test and lack of clinical contradictory evidence of the diagnosis (approximately 14 cases in Hiroshima).

Data Analysis. Contingency table analysis, using the appropriate chi-square tests as recommended by Fleiss, 12 was the principal method of analysis. All analyses were performed using only "confirmed diabetics" as diabetics. Relative risks were estimated by the corresponding odds ratios. The associated standard deviations were obtained as the products of the odds ratios and the square roots of the sums of the reciprocal cell counts. Relative risks were compared across groups using a Gaussian approximation. In the logistic regression analysis, four age categories and four total dose categories were used. Only in this analysis were the individuals required to have an estimated total dose. Average ages were computed over individuals who

陽性を示した対象者全員の確認を試みた. 少なくとも 135 mg/100 ml の空腹時血糖,又は180 mg/100 mlを 超える1時間及び2時間のOGTT値であれば、糖尿 病の診断の確認には十分であると考えた. また広島の 17症例及び長崎の4症例は診察周期以前に得た検査 結果を用いて確認した. これらの二つの基準はいずれ も本研究が行われた当時日本で用いられたもの11より も厳密であり、またこの集団に関する過去の研究2-4 で用いられたものよりも厳密である. 確認検査結果 のない対象者の医学記録については, 診断に臨床的 裏付け又は矛盾があるか検討した. 確認所見又は 矛盾する所見のない対象者は糖尿病とみなした. 長崎 では104症例を検討し、そのうち85症例は糖尿病の 臨床歴があり、19症例には矛盾する所見があった. 広島では244症例の検討を行い、そのうち146症例* には糖尿病の臨床歴,84症例*には矛盾する所見が あると推定されたが、14症例*には、どちらもないと 推定された.

要約すると、"確認された糖尿病"には、次のうちの少なくとも一つを有する患者をすべて含めた.

- a) 少なくとも2回, 血糖検査で陽性の結果を 示したが, 陽性検査を説明し得る既知の (非糖尿病) 状態を伴わない.
- b) 血糖検査で1回の陽性が認められ、糖尿病 の臨床歴がある.
- c) 血糖検査で1回の陽性が認められ,診断に 矛盾する臨床所見がない(広島で約14例).

データ解析.主な解析方法として、Fleiss¹²が推奨した適当なカイ二乗検定による分割表解析を用いた. "確認された糖尿病"のみを糖尿病とし、すべての解析を行った.相対危険度はそれに対応するオッズ比より推定した.それぞれの標準偏差は、オッズ比と各セル内の数の逆数の合計の平方根との積として得た. Gaussian 近似法を用いて群間で相対危険度を比較した.ロジスティック回帰解析では、四つの年齢群及び四つの総線量群を用いた.この解析においてのみ対象者の推定合計線量が必要であった。平均年齢はその診察周期に検診を受けた対象者全員から計算

^{*}The review records of 125 cases were lost. These estimates were obtained by extrapolating to the entire set of 244 cases the results of the 119 cases for which the review records still exist.

125 症例の検討記録が紛失していたため、これらの推定値は検討記録が残っている119 症例の結果を244の全症例に対して外挿することによって得た.

received an examination during that cycle. In this manner, censoring was accounted for. The across-group analyses of glycosuria and proteinuria rates were performed using the Mantel-Haenszel test assuming biennial urinalyses were independent. Diabetes incidence rates were computed for each cycle and city-sex group as the number of new cases diagnosed in the group during the cycle divided by the number of nonpreviously diagnosed examinees for that cycle and group. Diabetes prevalence rates were computed for each cycle and city-sex group as the number of diabetics examined divided by the total number of examinees.

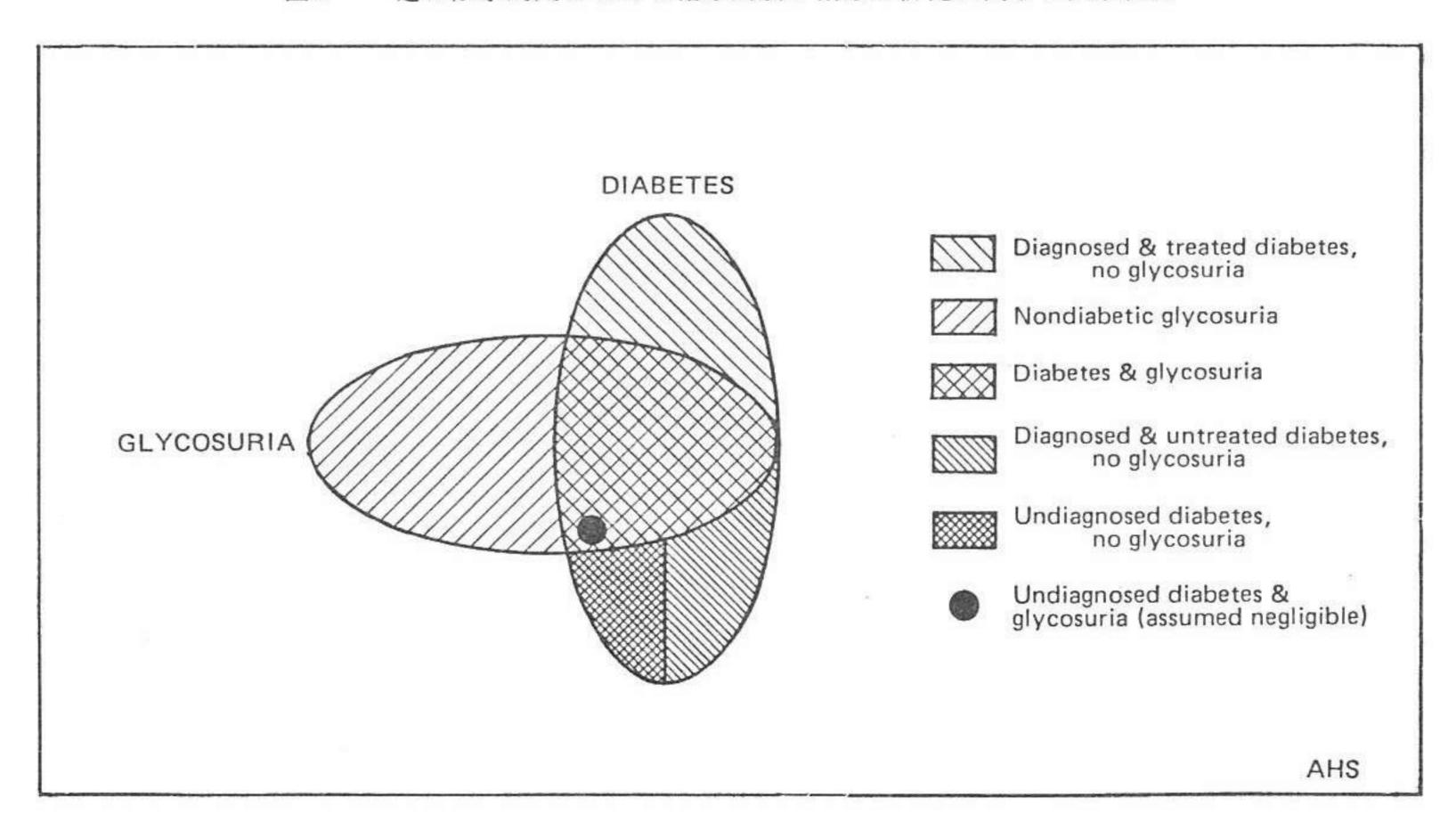
Since the initial screening test for diabetes was a routine test for glycosuria, and diabetics were not always glycosuric, the observed incidence and prevalence rates underestimated the actual morbidity rates. Readers who are familiar with Venn Diagrams may wish to review Figure 1, as it graphically demonstrates this underestimation problem. In all examination cycles most glycosurics who had not been previously diagnosed as diabetic were offered and received blood glucose tests. Thus, for each cycle, the largest group of undiagnosed diabetics were probably those who were not glycosuric at the time of their examination.

した. このようにして、センサリングが説明できる. 2年ごとの尿検査が独立したものであると仮定して、Mantel-Haenszel法を用いて糖尿率及び蛋白尿率の群間解析を行った. 各診察周期及び都市-性群における糖尿病発生率は、その周期中にその群内で新たに診断された症例数をその周期及び群内の以前に糖尿病の診断が下されていない被検者数で割って算出した. また、各診察周期及び都市-性群における糖尿病有病率は、糖尿病検査数を総被検者数で割って算出した.

最初の糖尿病スクリーニング検査は通常の糖尿検査であり、また、糖尿病患者に必ずしも糖尿が認められたわけではないので、観察発生率及び有病率は実際の罹病率よりも過小評価された。Venn Diagram に精通する読者は、この過小評価の問題を図示した図1の検討を希望するかもしれない。いずれの診察周期においても、糖尿を示したにもかかわらず過去に糖尿病と診断されたことのない者の大部分が血糖検査を受けた。したがって、各診察周期とも糖尿病診断が下されていない群のうち最大の部分は、恐らく検査時に糖尿を示さなかった者であろう。

FIGURE 1 SCHEMATIC REPRESENTATION OF THE CONCURRENCE OF DIABETES AND GLYCOSURIA FOR A GIVEN EXAMINATION CYCLE

図1 一定の診察周期における糖尿病及び糖尿の併発に関する図形表示



A statistical procedure was designed to adjust the observed diabetes incidence rate in each cycle and to overcompensate for the negative bias. It was hoped that the adjustment procedure would convey some idea of the extent of the negative biases in the observed incidence and prevalence rates. The adjustment procedure was based upon a probabilistic equality. By the definition of conditional probability,

各周期の観察糖尿病発生率を調整し、負の偏りを 過補償するように統計的手法を考案した. 調整手法 は、観察発生率及び観察有病率における負の偏りの 範囲についての考えを幾らか含むことが望ましかった. 調整手法は確率の等式を基にした. 条件つき確率の 定義によると、

Prob (diabetes and glycosuria) = Prob (glycosuria/diabetes) · Prob (diabetes) = Prob (diabetes/glycosuria) · Prob (glycosuria)

Hence, by division

したがって、除式にすると

$$Prob (diabetes) = \frac{Prob (diabetes/glycosuria) \cdot Prob (glycosuria)}{Prob (glycosuria/diabetes)}$$
(1)

For each cycle, the left-hand side of (1) is interpretable as the actual 2-year incidence rate of diabetes. If all glycosurics, without a prior diagnosis of diabetes, who refused a diabetes test were not diabetic, the numerator on the right-hand side of (1) would be estimated by the observed incidence. It is assumed that this is the case. Depending upon the extent to which this assumption is false, the observed incidence underestimates the numerator. The denominator is the adjustment to the observed incidence for the group of undiagnosed, nonglycosuric diabetics. If all diabetics were glycosuric, then (except for refusals) the observed incidence would equal the actual incidence because Prob (glycosuria/diabetes) would be unity. However, since diabetics were not always glycosuric, the observed incidence underestimated the actual incidence by about the proportion shown in (1).

For each cycle the incidence rates were adjusted using Hiroshima data on diabetic glycosuria after stratifying by age, sex, and presence or absence of obesity. The Prob (glycosuria/diabetes) values were estimated for all strata by using the glycosuria history of each diagnosed diabetic from their first observed cycle after diagnosis onward. As an attempt to intentionally underestimate Prob (glycosuria/diabetes) there was no consideration of whether a diabetic was under treatment at the time of urinalysis. It was assumed that within a stratum Prob (glyco-

各周期とも、(1) 式の左辺は2年間の糖尿病の実際の発生率であると解釈できる. 過去に糖尿病診断がないため糖尿病検査を拒否したが糖尿を示し、その人たちすべてが糖尿病でなかった場合、(1) 式の右辺の分子は観察発生率によって推定できるであろう. 実際、そういった場合もあると思われる. この仮定が誤りになる範囲で、観察発生率は分子を過小評価したものになる. 分母は、診断の下されていない非糖尿の糖尿病患者の観察発生率を調整したものである. 仮にすべての糖尿病患者が糖尿であるとすれば、Prob (糖尿/糖尿病) は1になるので、(診断拒否を除く) 観察発生率は実際の発生率に等しくなる. しかし、糖尿病患者が必ずしも糖尿を示すわけではないので、観察発生率は大体(1) 式に示した割合で実際の発生率を過小評価したものになった.

各診察周期の発生率は、糖尿を示す糖尿病患者に関する広島のデータを用い、年齢、性及び肥満症の有無別に層化した後、調整を行った、糖尿病の診断を下された対象者の診断後最初の観察周期以降の糖尿歴を用い、全層について Prob (糖尿/糖尿病)値を推定した、Prob (糖尿/糖尿病)を意図的に過小評価するために、尿検査時に糖尿病を治療中であったかどうかは考慮しなかった。層内では、Prob (糖尿/糖尿病)

suria/diabetes) did not change over time. Since glycosuria is a transient condition and the glycosuria measurements were taken about two years apart, they were considered statistically independent. Therefore, each diabetic yielded estimates, possibly in more than one stratum, of Prob (glycosuria/diabetes). The individual estimates were pooled into overall strata estimates by using the relative sample sizes as weights. Standard deviations of the stratum estimates were computed after considering extrabinomial variation. ¹³

There is a subtle issue concerning these estimates of Prob (glycosuria/diabetes) which requires further comment. They are based upon known diabetics, but they are intended to reflect the corresponding probabilities for undiagnosed diabetics. Therefore, it would appear that they are biased upwards. To counter this potential bias, the presence or absence of treatment for diabetes was ignored, as were the test results for the examination cycle when a diabetic had his/her first positive test. It is assumed that these two biases approximately canceled each other out. Of course, in the cycle of first positive test the glycosuria test was positive for almost all diabetics, since this was the main selection criterion used; it should thus not be used to calculate Prob (glycosuria/diabetes). However, ignoring this test result has an additional advantage. It ignores test results for the period when the diabetics would most generally be out of control and, thus, be least representative of the undiagnosed diabetics.

Prevalence estimates were obtained in a manner similar to that used for the incidence rates. In each stratum and examination cycle, observed prevalence was determined as the ratio of the number of diabetics examined during the cycle to the sample size of the stratum for that cycle. The adjusted prevalence was the sum of the number of examined diabetics who were diagnosed on a previous cycle and the number of new cases which corresponded to the adjusted incidence rate divided by the sample size of the stratum for that cycle. Suppose for a given cycle and stratum, 1,000 subjects were examined of which 50 had been previously diagnosed as diabetic and 10 were diagnosed on the current cycle. The observed incidence rate would be 10/950 and the adjusted incidence rate would be $(10/950) \times (1/0.5)$, if Prob (glycosuria/diabetes) was taken to be 0.5. The observed prevalence rate would be 60/1,000 and the adjusted prevalence rate would be [50+(10/0.5)]/1,000.

は経時変化しないものと仮定された。糖尿は一過的な 状態であり、また糖尿測定を約2年間隔で行ったので、 測定値は統計的に独立したものとみなした。したがっ て、各糖尿病患者に対し、恐らく一つ以上の層に おいて、Prob (糖尿/糖尿病)の推定値が得られた。 相対的な標本規模を加重として用い、個々の推定値 を全体の層の推定値にプールした。超過二項変動¹³ を考慮した後、層推定値の標準偏差を計算した。

Prob (糖尿/糖尿病)のこれらの推定値には、更に 説明を必要とする微妙な問題がある. これらの推定 値は既知の糖尿病例に基づいているが, それに対応 する診断の下されていない糖尿病患者の確率を反映 するよう意図されている. したがって, それらの値は 高値の方へ偏っているように見える. この潜在的な 偏りに対処するため、糖尿病患者が初回の陽性検査 を受けた診察周期の検査結果を無視したのと同様, 糖尿病の治療の有無を無視した. これら二つの偏りは ほぼ相殺したものと仮定する. もちろん, 最初の陽性 検査の診察周期では,ほとんどすべての糖尿病患者 が、糖尿検査で陽性を示した. それは、この検査結果 が主に使用された選択基準であるからである. した がって、それは Prob (糖尿/糖尿病)の計算には使用 すべきでない. しかし, この検査結果を無視すること にはもう一つの利点がある. これにより、糖尿病 患者例が全体的に最も照査されない期間, つまり, 診断の下されていない糖尿病患者を一番代表しない 期間の検査結果を無視することになるからである.

推定有病率は、発生率について用いたものと同じ 方法で求めた。各層及び各診察周期とも、観察有病 率はその周期の層の集団規模に対する、周期中に 検査を受けた糖尿病患者数の比として測定した。 調整有病率は、前周期に糖尿病と診断され検査を 受けた糖尿病患者数と、調整発生率に相当する新しい 症例の数を加え、その周期の合計であった層の集団 規模で割ったものとした。ある周期及び層において 1,000人の対象者が検査され、内50人は以前に糖尿 病と診断されたことがあり、10人は現在の周期で 糖尿病と診断されたと仮定する。観察発生率は、 10/950になり、Prob (糖尿/糖尿病)が0.5とすれ ば、調整発生率は(10/950)×(1/0.5)となる。 観察有病率は60/1,000となり、調整有病率は[50+ (10/0.5)]/1,000となる。

RESULTS

The number of AHS participants and the participation rates are given by cycle in Table 1. The participation rates are the percentages of those members of the cohort who were residing in the contacting area at the time of their scheduled biennial examination. There was very little difference in the participation rates by age or sex, but the Nagasaki rates were slightly higher than those in Hiroshima. In general the participation rates were quite high and the intercity differences relatively small.

結 果

AHS 対象者数及び受診率を周期別に表1に示す. 受診率は、2年ごとの診察時に連絡区域に居住していたコホート対象者の百分率である. 受診率には、年齢別又は性別による差がほとんどないが、長崎の受診率は広島よりやや高かった. 概して受診率は極めて高く、両市間の差異は比較的小さかった.

TABLE 1 NUMBER OF AHS PARTICIPANTS AND PERCENT PARTICIPATION BY CITY, SEX, AND CYCLE

表1 AHS 対象者数及び受診率: 都市, 性, 及び周期別

| | | Hiroshi | ima | Nagasaki | | | |
|-------|------------------|------------|-----------|-----------|-----------|--|--|
| Cycle | Year Examined | Male | Female | Male | Female | | |
| 1 | 1958-60 | 3495 88.6% | 6052 88.5 | 1242 90.4 | 1716 89.9 | | |
| 2 | 1960-62 | 3487 86.8 | 6199 89.0 | 1730 91.7 | 2443 92.1 | | |
| 3 | 1962-64 | 3267 85.3 | 5882 86.2 | 1763 92.0 | 2480 92.9 | | |
| 4 | 1964-66 | 3240 85.5 | 5871 86.4 | 1653 88.9 | 2392 89.7 | | |
| 5 | 1966-68 | 3038 84.0 | 5644 85.5 | 1549 86.1 | 2285 86.4 | | |
| 6 | 1968-70 | 2803 81.8 | 5384 83.3 | 1362 84.2 | 2030 85.6 | | |
| 7 | 1970-72 | 2584 79.1 | 4914 79.9 | 1392 82.2 | 2136 84.0 | | |
| 8 | 1972-74 | 2431 76.7 | 4708 77.6 | 1292 82.6 | 2000 83.0 | | |
| 9 | 1974-76 | 2166 72.1 | 4312 71.8 | 1274 80.0 | 1966 79.5 | | |
| 10 | 1976-78 | 1944 67.9 | 3777 66.3 | 1160 79.8 | 1805 78.7 | | |

Source: Minutes of the RERF Scientific Council, 28-29 September 1981

出所: 放影研専門評議員会議事録, 1981年9月28-29日

Since routine glycosuria determinations were the primary diabetes screening tests, a few introductory words on glycosuria in the AHS are appropriate. A detailed presentation of glycosuria in the AHS by age, city, and sex is given in Appendix Table 1. In general, for each city, glycosuria was more prevalent in males than in females (p<.001 in both cities) and for both sexes glycosuria was more prevalent in Hiroshima than in Nagasaki (p<.001 for both sexes). These differences by city and sex have been noted previously.^{3,4} Also, the expected increasing trends of rates due to aging were observed over 10 cycles.

The numbers of incident cases (i.e., newly confirmed diabetics) and blood glucose tests administered are shown by cycle in Table 2. In all analyses, only "confirmed diabetics" were considered to have noninsulin-dependent diabetes mellitus (NIDDM).

通常の糖尿測定は、第一次の糖尿病スクリーニング検査となったので、AHS内の糖尿について若干前置きが必要であろう。AHSにおける糖尿の年齢、都市、及び性別の詳細については付録表1に示した。概して、両市とも糖尿は女性よりも男性に多く(両市ともp<.001)、男女とも、糖尿は長崎よりも広島の方が多い(男女ともp<.001)。都市及び性別によるこれらの差異は以前にも認められている。3.4 また加齢による増加傾向の率も、10周期にわたって観察された。

発生症例数(すなわち,新しく確認された糖尿病患者数)及び血糖検査結果を周期別に表2に示す.全解析を通じて,"確認された糖尿病患者"のみがインシュリン非依存型糖尿病(NIDDM)であるとみなされた.

TABLE 2 NUMBER OF DIABETES SCREENING TESTS AND NUMBER OF DIAGNOSED DIABETICS

| =tr 0 | 姓 豆 兰 ュ 4 11 | 一、好松木粉 | ・1. 診断がよされた糖品庁庁間料 |
|-------|--------------|---------|-------------------|
| 表り | 想 旅 洒 スクリー | ーーノク検育報 | はと診断が下された糖尿病症例数 |

| | | | Mal | е | | Fen | nale |
|-------|--|-----|-----------------|---------------------|----------------------|---|---------|
| Cycle | Examined FBS OGTT Confirmed Diabetics/ | | No of Confirmed | (1) No of FBS | (1) No of OGTT | (2) No of Confirmed Diabetics/ Total First Time Positives | |
| | | | | Hiroshima | | | |
| 1 | 1958-60 | 274 | - | 73/92 | 212 | - | 56/62 |
| 2 | 1960-62 | 180 | 235 | 64/68 | 92 | 94 | 36/38 |
| 3 | 1962-64 | 60 | 263 | 17/18 | 74 | 178 | 19/19 |
| 4 | 1964-66 | 57 | 302 | 45/45 | 53 | 148 | 32/33 |
| 5 | 1966-68 | 23 | 317 | 36/37 | 18 | 236 | 50/53 |
| 6 | 1968-70 | 16 | 295 | 42/42 | 24 | 261 | 38/38 |
| 7 | 1970-72 | 17 | 209 | 31/31 | 53 | 128 | 29/30 |
| 8 | 1972-74 | 81 | 127 | 30/32 | 75 | 70 | 17/17 |
| 9 | 1974-76 | 81 | 149 | 26/27 | 77 | 63 | 25/25 |
| 10 | 1976-78 | 69 | 94 | 20/20 | 78 | 65 | 19/20 |
| | | | | 384/412 | | | 321/335 |
| | | | | Nagasaki | | | |
| 1 | 1958-60 | 35 | - | 7/8 | 24 | - | 2/2 |
| 2 | 1960-62 | 14 | 62 | 10/11 | 7 | 23 | 7/8 |
| 3 | 1962-64 | 20 | 100 | 4/9 | 21 | 64 | 10/10 |
| 4 | 1964-66 | 13 | 96 | 4/4 | 10 | 62 | 5/5 |
| 5 | 1966-68 | 20 | 166 | 11/14 | 25 | 150 | 10/12 |
| 6 | 1968-70 | 16 | 139 | 8/8 | 21 | 125 | 6/7 |
| 7 | 1970-72 | 3 | 16 | 1/1 | 6 | 14 | 2/2 |
| 8 | 1972-74 | 0 | 0 | 0/0 | 0 | 1 | 1/1 |
| 9 | 1974-76 | 0 | 0 | 0/0 | 0 | 0 | 0/0 |
| 10 | 1976-78 | 0 | 0 | 1/1 | 0 | 1 | 0/0 |
| | | | | 46/56 | | | 43/47 |

⁽¹⁾ Entries are the number of subjects who received tests for that cycle. 本欄はその周期に検査を受けた対象者数である.

The ratio of the total number of confirmed diabetics to the total number of subjects with at least one positive blood glucose test exceeded 93%. For both cities, the apparent^{1,14} male predominance among Japanese diabetics was again noted. The overall relative risks¹² (i.e., after collapsing over all cycles)

確認された糖尿病患者総数と、少なくとも1回の血糖検査で陽性を示した者の総数との比率は93%を超えた、両市とも、日本の糖尿病患者では明らかに男性が優勢であること^{1,14}が改めて確認された。男性の女性に対する糖尿病発生の総合的な相対危険度¹²(全周期を通じ下降後のもの)は、広島で2.1

⁽²⁾ Entries are of the form x/y, where x is the number of positive subjects who were later confirmed as diabetic, and y is the number of subjects who were positive but were not previously diagnosed as diabetic nor had this positive test be a confirmatory test. Consequently, the ratio x/y is the proportion of first-time positive cases on this cycle who were later confirmed as diabetic. 記載は x/y の形式で行った. この場合 x は陽性対象者で後に、糖尿病と確認された者の数であり、y は陽性ではあったが過去に糖尿病と診断されたことはなく、またその陽性と認められた検査が確認用検査にもならなかった対象者の数である. 結局、x/y の比は、この周期における初回の検査が陽性であり、後に糖尿病と確認された対象者の割合である.

of the occurrence of diabetes for males relative to females were 2.1 (±.18) in Hiroshima and 1.6 (±.35) in Nagasaki. The intercity difference was not significant (p=.23). For both sexes in Hiroshima there were unusually low incidence rates for cycle 3 (1962-64). Belsky et al⁴ suggested that there may have been more intensive screening for diabetes during cycles 1 (1958-60) and 2 (1960-62) than in several successive cycles. A possible result of this intensive screening program was that some diabetics, who might have been diagnosed during cycle 3 (1962-64) had the intensive screening not occurred, were actually diagnosed in cycle 2. Consequently, there may have been low incidence in cycle 3. In Nagasaki, beginning with cycle 7 (1970-72), there were few blood glucose tests performed. Therefore, all intercity comparisons of diabetes morbidity were restricted to the first six cycles (1958-70) only.

Previous investigations^{2-4,6,7} have not demonstrated any association between radiation and diabetes. In a preliminary analysis, the tendency was to confirm these prior findings by failing to find any relationships between radiation exposure and glycosuria or proteinuria. A subsequent logistic regression analysis of diabetes incidence and radiation exposure in Hiroshima confirmed this finding (Table 3, p>.50). Therefore, radiation exposure was not considered further in our analyses.

(土.18), 長崎で1.6(土.35)であった. 両市間の 差異は有意ではなかった(p=.23). 広島では男女 とも, 第3周期(1962-64年)の発生率が異常に低 かった. Belsky ら⁴ は, 第1周期(1958-60年)と 第2周期(1960-62年)には、それに続く数周期より も, 集中的なスクリーニングが行われた可能性がある ことを示唆した. この集中的なスクリーニング計画の 結果として考えられることは、糖尿病患者の何人かは、 実際には第2周期に診断されたが、もし集中的な スクリーニングが行われていなかったならば、第3周期 (1962-64年) になって診断されていたかもしれない ということである。第3周期の発生率が低かったのは そのためかもしれない. 長崎では第7周期(1970-72年)から、血糖検査はほとんど行われていない. したがって,糖尿病罹病率の両市間比較は,すべて 最初の6周期(1958-70年)に限定して行われた.

これまでの調査 $^{2-4,6,7}$ では、放射線と糖尿病の関係は明らかにされていない.予備的解析においては、放射線被曝と糖尿又は蛋白尿の関係を見いだすことができず、このような以前の結果を確認する傾向にあった.その後行われた広島の糖尿病発生率と放射線被曝に関するロジスティック回帰解析でも、この結果が確認されている(表3, p>.50).したがって今回の解析では、放射線被曝についてはもはや考慮しなかった.

TABLE 3 LOGISTIC REGRESSION OF OCCURRANCE OF DIABETES^a ON SEX, AGE, AND TOTAL DOSE IN HIROSHIMA

表3 広島における糖尿病の発生率 に関するロジスティック 回帰解析:性,年齢,及び総線量

| Covariate | Regression Estimate | SE | χ^2 (df=1) |
|-------------------------|---------------------|-------|-----------------|
| Intercept | -0.824 | 0.175 | 22.1 |
| Sex^b | -0.711 | 0.080 | 77.9 |
| Age ^c | -0.390 | 0.041 | 88.4 |
| Total dose ^d | 0.036 | 0.059 | 0.4 |

- a Based upon confirmed diabetics excluding all subjects without estimated total dose.
 - 推定総線量のない対象者全員を除く確認された糖尿病患者に基づく.
- b Categorized as 1=Male and 2=Female. 1=男性, 2=女性に分類.
- c Categorized on birth year as 1=Before 1901, 2=1901-15, 3=1916-30, and 4=1931-45.
- 出生年を1=1901年以前,2=1901-15年,3=1916-30年,4=1931-45年 に分類. d Categorized on estimated total dose as 1=No exposure, 2=1-200 rad,
- 3=201-400 rad, and 4=401-600 rad. 推定総線量を1=非被曝, 2=1-200 rad, 3=201-400 rad, 及び4=401-600 rad に分類.

Table 4 contains the average weights and ages of the diabetic and all other subjects. Diabetics were generally older than those subjects without known diabetes, but for the four city-sex groups the average ages of the diabetics were about the same. The average age at diagnosis was slightly higher in Hiroshima (males: 57.1 ± 0.7 , females: 57.7 ± 0.9) than in Nagasaki (males: 54.8 ± 1.1 , females: 53.8 ± 1.2). For both cities the male diabetics were nonobese, with average body weights for all 10 cycles within the 100% to 108% IBW range. The female diabetics tended to be slightly heavier than the male diabetics with average weights usually in the 109% to 114% IBW range. At all examinations from diagnosis onward, 28.9% of the diabetic males were obese and 43.6% of the diabetic females were obese.

表 4 には、糖尿病患者及びその他の対象者全員の 平均体重及び平均年齢を示す. 糖尿病患者は一般に 糖尿病が認められていない者よりも年齢が高かった が、都市-性の4群については、糖尿病患者の平均 年齢はほぼ等しかった. 診断時の平均年齢は, 広島 (男性: 57.1±0.7, 女性: 57.7±0.9)の方が長崎 (男性: 54.8±1.1, 女性: 53.8±1.2)よりもわずか に高かった. 両市とも, 男性の糖尿病患者は, 10周期 全体を通じて平均体重が100%から108% IBW の範囲 にあり、肥満を伴わなかった. 女性の糖尿病患者は, 平均体重が大体109%から114% IBW の範囲であり, 男性の糖尿病患者よりもわずかに重い傾向にあった. 診断後の全診察では、男性の糖尿病患者の28.9%が 肥満であり、女性では43.6%であった.

TABLE 4 AGES AND IDEAL BODY WEIGHTS (IBW) OF DIABETICS AND ALL OTHER SUBJECTS^a 糖尿病患者及びその他すべての対象者の年齢と理想体重(IBW)

| | | | | | Cy | cle | | | | |
|-----------------------------------|-------|-------|-------|----------|-------|-------|-------|-------|-------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | H | iroshima | Male | | | | | |
| Average Age | | | | | | | | | | |
| All diabetics | 55.7 | 57.5 | 58.8 | 58.7 | 59.3 | 60.8 | 61.2 | 61.1 | 62.9 | 63.3 |
| Diabetics diagnosed on this cycle | 55.7 | 56.3 | 56.4 | 55.3 | 58.8 | 60.9 | 58.8 | 54.0 | 61.3 | 59.9 |
| All others | 45.0 | 46.8 | 48.4 | 49.8 | 51.2 | 52.4 | 53.6 | 54.8 | 55.9 | 57.2 |
| Average IBW | | | | | | | | | | |
| All diabetics | 102.7 | 102.9 | 102.1 | 104.4 | 105.3 | 104.5 | 106.1 | 106.3 | 106.6 | 106.9 |
| Diabetics diagnosed on this cycle | 102.7 | 102.8 | 109.1 | 109.4 | 107.3 | 106.1 | 108.0 | 111.9 | 108.6 | 113.1 |
| All others | 97.6 | 97.8 | 97.6 | 98.5 | 99.5 | 100.6 | 101.3 | 101.2 | 102.1 | 102.8 |
| | | | Hir | oshima F | emale | | | | | |
| Average Age | | | | | | | | | | |
| All diabetics | 55.9 | 57.7 | 58.0 | 58.9 | 58.6 | 60.5 | 61.4 | 62.7 | 63.2 | 63.1 |
| Diabetics diagnosed on this cycle | 55.9 | 58.1 | 55.3 | 57.3 | 54.7 | 61.0 | 60.3 | 58.1 | 59.8 | 60.2 |
| All others | 43.7 | 45.6 | 47.4 | 49.0 | 50.7 | 52.2 | 53.9 | 55.2 | 56.3 | 57.7 |
| Average IBW | | | | | | | | | | |
| All diabetics | 109.4 | 109.3 | 109.6 | 110.5 | 113.8 | 113.3 | 113.4 | 112.2 | 112.4 | 112.3 |
| Diabetics diagnosed on this cycle | 109.4 | 111.5 | 113.3 | 110.1 | 119.2 | 114.6 | 120.0 | 116.8 | 116.0 | 114.0 |
| All others | 100.4 | 100.8 | 100.9 | 101.9 | 103.1 | 104.1 | 104.9 | 105.2 | 106.0 | 106.3 |

TABLE 4 (Continued 続き)

| | | | | | C | ycle | | | | |
|--------------------------------------|--------------------|-------|-------|-----------|-------|-------|--------------------|-------------|-------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | |] | Nagasaki | Male | | | | | |
| Average Age | | | | | | | | | | |
| All diabetics | 58.6 | 55.2 | 54.2 | 56.3 | 57.5 | 58.6 | 58.9 | 64.5 | 63.7 | 63.3 |
| Diabetics diagnosed on this cycle | 58.6 | 53.4 | 51.3 | 55.3 | 55.5 | 55.9 | 39.0 ^b | | | |
| All others | 39.5 | 42.1 | 44.0 | 45.9 | 47.2 | 48.8 | 50.0 | 51.9 | 53.3 | 54.9 |
| Average IBW | | | | | | | | | | |
| All diabetics | 99.9 | 103.8 | 105.4 | 104.4 | 106.7 | 105.5 | 103.9 | 107.7 | 107.6 | 106.0 |
| Diabetics diagnosed on this cycle | 99.9 | 106.1 | 102.8 | 104.3 | 107.7 | 109.9 | 112.0 ^b | ~ | _ | _ |
| All others | 98.1 | 97.5 | 98.2 | 99.3 | 99.6 | 100.4 | 101.7 | 101.9 | 101.9 | 103.0 |
| | | | N | agasaki F | emale | | | | | |
| Average Age | | | | J | | | | | | |
| All diabetics | 51.5° | 56.4 | 57.0 | 55.1 | 55.1 | 56.1 | 58.0 | 57.9 | 61.0 | 60.3 |
| Diabetics diagnosed | | | | | | | | | | |
| on this cycle | 51.5° | 57.3 | 56.0 | 45.2 | 53.6 | 55.2 | 55.5 ^c | 45.0^{b} | - | - |
| All others | 37.5 | 39.3 | 41.3 | 43.0 | 44.6 | 46.2 | 48.3 | 49.9 | 51.6 | 53.2 |
| Average IBW | | | | | | | | | | |
| All diabetics | 113.5 ^c | 112.3 | 108.9 | 106.8 | 113.6 | 109.1 | 111.1 | 107.7 | 105.6 | 102.0 |
| Diabetics diagnosed on this cycle | 113.5° | 112.7 | 112.6 | 109.5 | 122.2 | 97.8 | 114.5° | 108.0^{b} | _ | _ |
| All others | 99.6 | 99.6 | 100.4 | 101.7 | 102.6 | 103.9 | 105.1 | 106.1 | 106.3 | 107.0 |

a All other subjects are all subjects without known diabetes.

Table 5 shows the relative risks for diabetes due to a positive family history. In Hiroshima the relative risks for both sexes were between 2.5 and 3.5, with no significant differences between males and females (p>.50). In Nagasaki the numbers were too small to make definite statements, though it appears that females were not at an increased risk due to a positive family history. Regarding intercity differences, 4.1% of Hiroshima subjects reported a positive family history, whereas in Nagasaki the reported rate was only 2.1%. The differences were highly significant (males: p=.001; females: p<.001; both sexes combined: p<.001) and they reflect the higher rates of diabetes in Hiroshima. Hiroshima females reported a higher rate of family history than did Hiroshima males (p=.005), but in Nagasaki there was no difference between the sexes (p>.50).

表 5 には、陽性の家族歴による糖尿病の相対危険度を示す。相対危険度は、広島では男女間に有意差がなく(p>.50)、男女とも 2.5から 3.5の間であった。長崎では、明確な所見を述べるにはあまりにも少数であったが、陽性の家族歴による危険度の増加は女性にはないと思われた。両市間の差異については、広島の対象者の 4.1% は陽性の家族歴を報告しているが、長崎ではわずかに 2.1% であった。この差は極めて有意であり (男性: p=.001; 女性: p<.001; 男女: p<.001)、これは広島の糖尿病の高い率を反映するものである。広島の女性は家族歴について男性よりも高率を示したが (p=.005)、長崎では男女間に差がなかった (p>.50).

その他すべての対象者とは, 既知の糖尿病が認められないすべての対象者である.

b Based upon one case. 1症例に基づく.

c Based upon two cases. 2 症例に基づく.

| TABLE 5 | FAMILY HISTORY OF DIABETES | |
|---------|----------------------------|--|
| | 表 5 糖尿病の家族歴 | |

| | Diabetes | No Diabetes | Total | Relative Risk of Diabetes Due to a Positive Family History ^a |
|------------|----------|----------------|----------|--|
| | | Hiro | shima M | ale |
| History | 32 | 152 | 184 | |
| No History | 352 | 4727 | 5066 | 2.8 ± .6 |
| | | Hiros | hima Fer | male |
| History | 41 | 340 | 380 | |
| No History | 280 | 7814 | 8088 | $3.4 \pm .6$ |
| | | Nag | asaki Ma | ale |
| History | 3 | 54 | 57 | |
| No History | 43 | 2616 | 2659 | 3.4 ± 2.1 |
| | | Naga | saki Fen | nale |
| History | 1 | 75 | 76 | 10422 (124 125 122 122 122 122 122 122 122 122 122 |
| No History | 42 | 3419 | 3461 | 1.1 ± 1.1 |

a Estimated relative risk ± 1 SD of estimate. 推定相対危険度±推定値の1SD.

The possible relationships between a positive family history, sex, and glucose intolerance were examined by considering average OGTT curves. The means and standard errors of the 2-hour values for three typical cycles are given in Table 6. The comparisons for the 1-hour values give results virtually identical to those shown. Even though many of the standard errors are relatively large, and thus there are few significant differences, there is a suggestive pattern. The post-loading values were evaluated in females when compared to males, as was reported also by Belsky et al. 15 Moreover, for both sexes the OGTT curves of individuals with positive family histories of diabetes were elevated relative to those of individuals without family histories. One might speculate, as suggested by a reviewer, that the higher average blood glucose level associated with positive family history is due in part to more diabetics in that group, as shown in Table 5.

Table 7 contains the glycosuria and proteinuria prevalence rates for the 10 cycles stratified by city, sex, and presence or absence of known diabetes. For all four city-sex groups, a trace to 4+ glycosuria (p<.001 in all cases) and 1+ or greater proteinuria (p<.001 in Hiroshima and p≤.005 in Nagasaki) were more common in diabetics than in all other subjects. As noted above, for both sexes glycosuria was more prevalent in Hiroshima than in Nagasaki

陽性の家族歴、性、及びブドウ糖不耐性の間に考えられる関係については、平均 OGTT 曲線を検討して調べた。典型的な三つの周期に関する 2 時間値の平均及び標準誤差を表 6 に示す。 1 時間値に関する比較は、そこに示す値と大体同じ結果を示した。標準誤差の多くが相対的に大きく、したがって有意差はほとんどないが、示唆的なパターンは認められる。Belsky 5¹⁵ も報告したように、男性と比較するときには、女性の場合、後に加重した値を評価した。更に、糖尿病の家族歴をもつ各対象者の OGTT 曲線は、男女とも、家族歴をもたない者に比べ高かった。本稿の検討者が示唆するように、陽性の家族歴に関連して平均血糖値が高いということは、表 5 に示すとおり、その群に糖尿病患者が多いことに幾らか原因があると推測できるかもしれない。

表7は、10周期における糖尿及び蛋白尿の有病率を都市別、性別、及び既知の糖尿病の有無により層化したものである.都市-性分類の4群全体においては、極微量から4十の糖尿(全症例でp<.001)及び1+以上の蛋白尿(広島でp<.001及び長崎でp≤.005)は、ほかの全対象者よりも糖尿病患者に多かった.上記のとおり男女とも糖尿は広島の方が長崎より

TABLE 6 AVERAGE 2-HOUR OGTT VALUES (mg/100 ml) FOR THREE REPRESENTATIVE CYCLES: HIROSHIMA

表 6 代表的な 3 周期についての平均 OGTT 2 時間値 (mg/100 ml): 広島

| | Family History of Diabetes | | Cycle 4 | Cycle 7 | Cycle 9 |
|--------|-------------------------------|-------------------------|---------|---------|---------|
| Male | Yes | $\overline{\mathbf{X}}$ | 171.3 | 199.2 | 188.6 |
| | | (SE) | (10.5) | (20.7) | (19.8) |
| | No | $\overline{\mathbf{X}}$ | 159.9 | 176.8 | 169.7 |
| | | (SE) | (4.5) | (6.4) | (7.0) |
| Female | Yes | $\overline{\mathbf{X}}$ | 204.8 | 255.2 | 256.7 |
| | | (SE) | (28.5) | (44.1) | (47.8) |
| | No | $\overline{\mathbf{X}}$ | 186.3 | 204.5 | 210.6 |
| | | (SE) | (9.3) | (9.6) | (12.1) |

TABLE 7 PREVALENCE RATES (CASES/100) OF GLYCOSURIA a,b AND PROTEINURIA c

表 7 糖尿 a,b 及び蛋白尿 c の有病率 (症例数/100)

| | | | | | Cyc | le | | | | |
|------------------------|------|------|------|----------|---------|------|------|------|------|------|
|). | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | I | Iiroshin | na Male | | | | | |
| Glycosuria | | | | | | | | | | |
| Diabetics | - 1 | 65.5 | 66.7 | 51.7 | 71.6 | 54.2 | 51.2 | 31.3 | 58.0 | 50.3 |
| All others | 7.5 | 7.0 | 7.9 | 8.0 | 11.8 | 8.7 | 7.8 | 4.0 | 7.1 | 5.6 |
| Total | 9.2 | 9.4 | 10.3 | 10.6 | 15.3 | 11.9 | 11.2 | 6.5 | 11.9 | 10.2 |
| Proteinuria | | | | | | | | | | |
| Diabetics ^d | 21.9 | 14.2 | 14.2 | 24.4 | 27.5 | 21.3 | 22.2 | 27.8 | 23.3 | 28.0 |
| All others | 6.5 | 6.3 | 5.0 | 3.7 | 3.9 | 3.9 | 4.0 | 3.3 | 3.8 | 6.0 |
| Total | 7.1 | 6.5 | 5.6 | 4.7 | 5.3 | 5.1 | 5.3 | 5.2 | 5.4 | 7.9 |
| | | | H | iroshima | a Femal | e | | | | |
| Glycosuria | | | | | | | | | | |
| Diabetics | - | 69.4 | 67.6 | 60.9 | 61.4 | 61.1 | 51.4 | 38.1 | 52.4 | 47.8 |
| All others | 3.4 | 1.4 | 1.5 | 1.8 | 2.8 | 2.2 | 1.5 | 1.2 | 1.5 | 1.7 |
| Total | 4.2 | 2.5 | 2.5 | 3.0 | 4.4 | 4.2 | 3.4 | 2.6 | 3.7 | 3.6 |
| Proteinuria | | | | | | | | | | |
| Diabetics ^d | 23.2 | 25.3 | 23.6 | 24.2 | 18.9 | 17.1 | 22.0 | 27.7 | 29.2 | 28.2 |
| All others | 5.8 | 5.5 | 4.6 | 4.0 | 3.8 | 3.6 | 3.8 | 3.6 | 3.8 | 5.6 |
| Total | 6.0 | 5.8 | 5.0 | 4.4 | 4.2 | 4.0 | 4.4 | 4.6 | 4.8 | 6.5 |

TABLE 7 (Continued 続き)

| | | | | | Cyc | le | | | | |
|-------------|-------------------|-------------|------------|----------|---------|------|------|------|------|------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | 92. 9.4 | Nagasal | ci Male | | | | | |
| Glycosuria | | | | | | | | | | |
| Diabetics | () | 75.0 | 75.0 | 75.0 | 58.8 | 70.6 | 48.2 | 34.8 | 57.9 | 65.0 |
| All others | 3.9 | 3.8 | 6.0 | 4.7 | 8.3 | 9.9 | 5.3 | 3.6 | 4.3 | 6.1 |
| Total | 4.4 | 4.5 | 6.6 | 5.4 | 9.5 | 11.0 | 6.2 | 4.1 | 5.1 | 7. |
| Proteinuria | | | | | | | | | | |
| Diabetics | 14.3 | 6.7 | 17.7 | 6.3 | 17.9 | 12.0 | 17.9 | 26.1 | 36.8 | 19. |
| All others | 4.8 | 2.7 | 3.0 | 3.1 | 3.1 | 4.2 | 4.5 | 5.3 | 7.1 | 5.9 |
| Total | 4.8 | 2.8 | 3.1 | 3.2 | 3.4 | 4.3 | 4.7 | 5.7 | 7.5 | 6. |
| | | | N | lagasaki | Female | ; | | | | |
| Glycosuria | | | | | | | | | | |
| Diabetics | 73 - 1 | 100.0^{e} | 75.0 | 75.0 | 63.2 | 57.7 | 45.2 | 29.0 | 46.4 | 45.4 |
| All others | 2.0 | 0.7 | 1.1 | 1.7 | 1.9 | 3.6 | 1.0 | 0.8 | 1.6 | 1.3 |
| Total | 2.7 | 1.0 | 1.6 | 2.3 | 2.6 | 4.4 | 1.6 | 1.2 | 2.2 | 2.: |
| Proteinuria | | | | | | | | | | |
| Diabetics | 0^e | 22.2 | 16.7 | 9.5 | 10.3 | 12.5 | 23.5 | 33.3 | 24.1 | 17. |
| All others | 5.7 | 2.1 | 2.3 | 3.0 | 3.5 | 3.7 | 3.4 | 3.6 | 5.1 | 4.9 |
| Total | 5.7 | 2.2 | 2.4 | 3.1 | 3.6 | 3.8 | 3.7 | 4.1 | 5.4 | 5.0 |

a Glycosuria is trace or greater. 糖尿は極微量又はそれ以上.

and in both cities glycosuria was more prevalent in males than in females. Among those subjects without known diabetes, the glycosuria rates were higher (p<.001 for both sexes) in Hiroshima than in Nagasaki. However, among diabetics, the glycosuria rates were about the same in all four city-sex groups. Thus, some of the intercity differences in diabetes morbidity may be attributed to the combination of routine urinalysis for diabetes prescreening and the higher glycosuria rates in Hiroshima. It seems doubtful, however, if this would explain all of the intercity differences in diabetes morbidity. In Hiroshima, glycosuria was about three times more prevalent in males than in females (p<.001), but diabetes was about 1.4 times more prevalent in glycosuric females (45%) than in glycosuric males (32%). In several populations, including 多発であった. 既知の糖尿病が認められない対象者では、糖尿率は広島の方が長崎よりも高かった (男女とも p<.001). しかし、糖尿病患者においては糖尿率は都市-性の 4 群すべてにおいてほぼ等しかった. したがって、糖尿病罹病率における両市間の差異の幾らかは、糖尿病の前スクリーニングのための通常の尿検査と、広島の高い糖尿率の組み合わせによるものかもしれない. しかし、このことが糖尿病罹病率における両市間の差異をすべて説明するかどうかは疑わしいように思われる. 広島では、糖尿の有病率は、男性の方が女性よりも約3倍高い (p<.001) が、糖尿病の有病率は、糖尿を示す女性 (45%) の方が糖尿病の有病率は、糖尿を示す女性 (45%) の方が

b Excluding diabetics during diagnosis cycle. 診断周期中の糖尿病を除く.

c Proteinuria is 1+ or greater. 蛋白尿は1+又はそれ以上.

d Excluding 13 male and 7 female diabetics who were diagnosed after cycle 6 but whose proteinuria values were unknown.

第6周期以後に診断されたが、蛋白尿値が不明であった男性13例、及び女性7例の糖尿病患者を除く、

e Based on 2 cases. 2症例に基づく.

Japanese, it has been shown¹ that males have a higher rate of glycosuria than do females, but among glycosurics, males have a lower rate of diabetes than do females. For both sexes, proteinuria was more common in Hiroshima than in Nagasaki, but the small differences are probably attributable to the higher prevalence of diabetes in Hiroshima. Among those subjects without known diabetes, the rates of proteinuria were about the same in all four citysex groups (p>.10), and the average prevalence was 4.3%. Among diabetics, within each city the proteinuria rates were about the same for both sexes (p>.25 for both cities), though the Nagasaki numbers were small. The prevalence of proteinuria in Hiroshima diabetics was approximately 25.6% and in Nagasaki diabetics it was approximately 18.7%. A detailed presentation of glycosuria and proteinuria in the AHS by age, city, and sex is given in Appendix Table 1.

The estimated probabilities used in equation (1) for adjusting the observed incidence and prevalence rates are given in Table 8. Within the age- and sexstrata there was little effect (p>.50 for males; p>.10 for females) on the estimates due to the presence or absence of obesity. Therefore, the estimates within strata were pooled over weight groups and these pooled estimates were used for the adjustment. Because the glycosuria rates among diabetics of the same sex in the two cities were quite similar and the Nagasaki numbers were small, the estimates of Prob (glycosulia/diabetes) from Hiroshima were used to adjust the Nagasaki rates. The observed and adjusted incidence and prevalence rates for diabetes are given in Table 9. A more detailed presentation by age is given in Appendix Table 2. In Hiroshima, the observed incidence for males from cycle 4 (1964-66) onward was approximately constant (p>.50) at about 1.2-1.3 cases/100 per cycle, while for the females (except for cycle 5) the observed incidence from cycle 4 onward was approximately constant (p>.20) at 0.5-0.6 cases/100 per cycle.

Since the adjustments of incidence and prevalence rates were solely intended to indicate the extent of negative biases in the observed rates and because the accuracy and precision of the adjusted rates are unknown, all analyses of incidence and prevalence rates, unless otherwise noted, refer to the observed rates. 男性 (32%) よりも約1.4倍高かった. 日本人を含む 幾つかの集団においては、男性は女性よりも糖尿率 が高いが、糖尿を示す者の中では、男性の方が女性 よりも糖尿病率が低いことが認められている.1 男女 とも、蛋白尿は長崎よりも広島の方に多かったが、 この小さな差は、恐らく広島の糖尿病有病率が高い ことによるものであろう. 既知の糖尿病が認められ ない対象者においては、蛋白尿率は都市-性の4群 すべてにおいてほぼ等しく(p>.10)、平均有病率は 4.3%であった. 糖尿病患者においては、長崎の症例 数は少なかったが、蛋白尿率は両市とも男女ほぼ同じ であった(両市ともp>.25). 広島の糖尿病患者にお ける蛋白尿有病率は約25.6%であり、長崎の場合は 約18.7%であった. 付録表1に、AHS 内の糖尿及び 蛋白尿率の詳細を、年齢、都市、及び性別に示した.

観察発生率及び観察有病率を調整するために方程式(1) で用いた確率推定値を表8に示す. 年齢及び性の 層内では,推定値が肥満症の有無によって影響 されることはほとんどなかった(男性: p>.50; 女性: p>.10). したがって層内の推定値を体重群 にプールし、このプールした推定値を調整に用いた. 両市とも,同性の糖尿病患者における糖尿率はかな り近似しており、また長崎の数は少なかったので、 広島の Prob (糖尿/糖尿病)の推定値を用いて長崎 の率の調整を行った. 糖尿病に関する観察及び調整 発生率並びに有病率を表9に示す.年齢別の詳細 は付録表2に示した. 広島では, 第4周期(1964-66年)以降の男性の観察発生率は、1周期当たり 約1.2-1.3症例/100でほぼ一定しており(p>.50), 第4周期以降(第5周期を除く)の女性の観察発生 率は,1周期当たり約0.5-0.6症例/100でほぼ一定 している(p>.20).

発生率及び有病率の調整は単に観察発生率における負の偏りの範囲を示すためだけに行ったものであり、また、調整率の正確さ及び精度は明らかでないので、発生率と有病率の解析はすべて、特に記述がなければ、観察発生率に関するものである.

TABLE 8 ESTIMATED PROBABILITIES OF DIABETIC GLYCOSURIA 表 8 糖尿を示す糖尿病患者の確率推定値

| | Wai 14 C | Probabilities (SE) | | | | | | | |
|---------|--------------|---------------------------|-------------|------------------|-------------|--|--|--|--|
| Age | Weight Group | Hiroshin | ma Male | Hiroshima Female | | | | | |
| <40 | Nonobese | .400 (.097) | 210 (0(0) | .400 (.116) | 244 (007) | | | | |
| | Obese | .333 (.149) | .319 (.069) | .500 (.183) | .344 (.087) | | | | |
| 40-49 | Nonobese | .466 (.068) | (.068) | .417 (.075) | 121 (060) | | | | |
| | Obese | e .605 (.101) .514 (.056) | .514 (.056) | .432 (.102) | .424 (.060) | | | | |
| 50-59 | Nonobese | .481 (.054) | 404 (044) | .514 (.058) | 450 (041) | | | | |
| | Obese | .526 (.082) | .494 (.044) | .390 (.057) | .452 (.041) | | | | |
| 60+ | Nonobese | .465 (.028) | 449 (025) | .528 (.036) | 490 (020) | | | | |
| | Obese | .368 (.058) | .448 (.025) | .372 (.048) | .480 (.029) | | | | |
| Overall | | | .445 (.020) | | .462 (.021) | | | | |

Standard errors are given in parentheses. 括弧内は標準誤差.

TABLE 9 OBSERVED AND ADJUSTED INCIDENCE AND PREVALENCE RATES (CASES/100) FOR DIABETES

表 9 糖尿病の観察及び調整発生率並びに有病率(症例/100)

| | | | | | (| Cycle | | | | |
|---------------------|---------------|-----|--------|---------|-------------------|-------|-----|-----|------|------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | Hiros | hima l | Male | | | | | |
| Observed Incidence | - | 1.9 | .5 | 1.4 | 1.2 | 1.6 | 1.2 | 1.3 | 1.3 | 1.1 |
| Adjusted Incidence | - | 4.0 | 1.2 | 3.1 | 2.7 | 3.4 | 2.7 | 2.9 | 2.9 | 2.5 |
| Observed Prevalence | 2.1 | 3.4 | 3.8 | 5.0 | 5.6 | 6.8 | 7.7 | 8.3 | 8.8 | 9.7 |
| Adjusted Prevalence | - | 5.6 | 4.4 | 6.6 | 7.0 | 8.6 | 9.1 | 9.7 | 10.2 | 10.0 |
| | | | Hirosh | ima F | emale | | | | | |
| Observed Incidence | _ | .6 | .3 | .6 | .9 | .7 | .5 | .4 | .6 | .5 |
| Adjusted Incidence | - | 1.2 | .7 | 1.2 | 1.9 | 1.6 | 1.2 | .8 | 1.3 | 1.1 |
| Observed Prevalence | .9 | 1.4 | 1.5 | 2.0 | 2.7 | 3.4 | 3.5 | 3.8 | 4.1 | 4.1 |
| Adjusted Prevalence | - | 2.0 | 1.9 | 2.7 | 3.7 | 4.2 | 4.1 | 4.2 | 4.7 | 4.7 |
| | | | Naga | saki M | ale^a | | | | | |
| Observed Incidence | - | .6 | .2 | .2 | .7 | .6 | .1 | - | - | .09 |
| Adjusted Incidence | - | 1.2 | .5 | .5 | 1.6 | 1.3 | .2 | - | - | .2 |
| Observed Prevalence | .6 | .8 | 1.0 | 1.0 | 1.8 | 1.8 | 2.0 | 1.8 | 1.5 | 1.8 |
| Adjusted Prevalence | - | 1.5 | 1.2 | 1.2 | 2.6 | 2.5 | 2.1 | 1.8 | 1.5 | 1.9 |
| | | | Nagas | aki Fer | nale^a | | | | | |
| Observed Incidence | - | .3 | .4 | .2 | .4 | .3 | .1 | .1 | - | 3 |
| Adjusted Incidence | - | .6 | .9 | .4 | .9 | .6 | .2 | .1 | - | Ĩ |
| Observed Prevalence | .1 | .4 | .7 | .9 | 1.3 | 1.6 | 1.6 | 1.6 | 1.5 | 1.3 |
| Adjusted Prevalence | . | .7 | 1.2 | 1:1 | 1.8 | 1.9 | 1.7 | 1.7 | 1.5 | 1.3 |

a Adjusted using the sex-specific estimates in Table 8. 表 8 の性別推定値を用いて調整したもの.

Figures 2 and 3 show the age-specific and overall observed and adjusted prevalence rates for the Hiroshima males and females. The overall prevalence increased in time (males: 0.75 ± 0.04 cases/100 per cycle; females: 0.36 ± 0.04 cases/100 per cycle) and the increase appeared due to individuals aged 50 or more. It should be emphasized that over the 20 years of study, the age distributions of the cohort have changed. For example, during cycle 1 (1958-60), 23.5% of the Hiroshima males were aged 60 or more, while at cycle 10 (1976-78) 45.7% of the Hiroshima males were aged 60 or more.

One explanation for the increasing prevalence rates is that they are due to constant age-specific incidence rates and an aging population. Since NIDDM is not a rapidly fatal disease, the number of surviving diabetics might increase for a long period of time. If this were the case, as the population aged the prevalence rate would increase. Note, however, there may be other explanations as well.

DISCUSSION

Findings regarding diabetes, glycosuria, and proteinuria in a 20-year cohort study of Japanese residents of Hiroshima and Nagasaki were presented. Surprisingly, there have been few reports of any previous long-term population studies of diabetes among the Japanese. Previous short-term studies^{2,3} of the AHS cohort found diabetes and glycosuria were more prevalent in Hiroshima than in Nagasaki, and in both cities more prevalent in males than in females. Also, persons in Hiroshima, irrespective of sex, were reported to be at increased risk for diabetes among those with a positive family history of diabetes. However, there was some disagreement^{3,4} regarding increased proteinuria rates among diabetics. Recently, the NDDG¹⁰ proposed more stringent criteria for diabetes than that used in these previous studies. Therefore, the present study was undertaken to see how the previous findings might change when the NDDG criteria was used and to update the findings regarding diabetes in the AHS through 1978.

There are two pertinent facts about diabetes in the AHS worth noting. First, IDDM is quite rare in Japan when compared to the West. In fact, in the entire AHS cohort of approximately 20,000 subjects there were no cases of IDDM. Second, there were few blood glucose tests performed in Nagasaki after cycle 6 (1968-70). Therefore, intercity comparisons んど血糖検査が行われていないことである. したがっ

図2及び3は、広島の男女に関する年齢別及び全体 的な観察有病率並びに調整有病率を示す. 全体的な 有病率は経時的に増加したが(男性: 1周期当たり 0.75 ± 0.04 症例 / 100; 女性: 1 周期当たり 0.36 ± 0.04症例/100),この増加は50歳以上の者によって もたらされているようであった. コホートの年齢分布 が20年にわたる研究の間に変化していることを強調 しておかねばならない. 例えば, 第1周期(1958-60年) には、広島の男性のうち60歳以上の者は23.5% であったが、第10周期(1976-78年)には45.7%で あった.

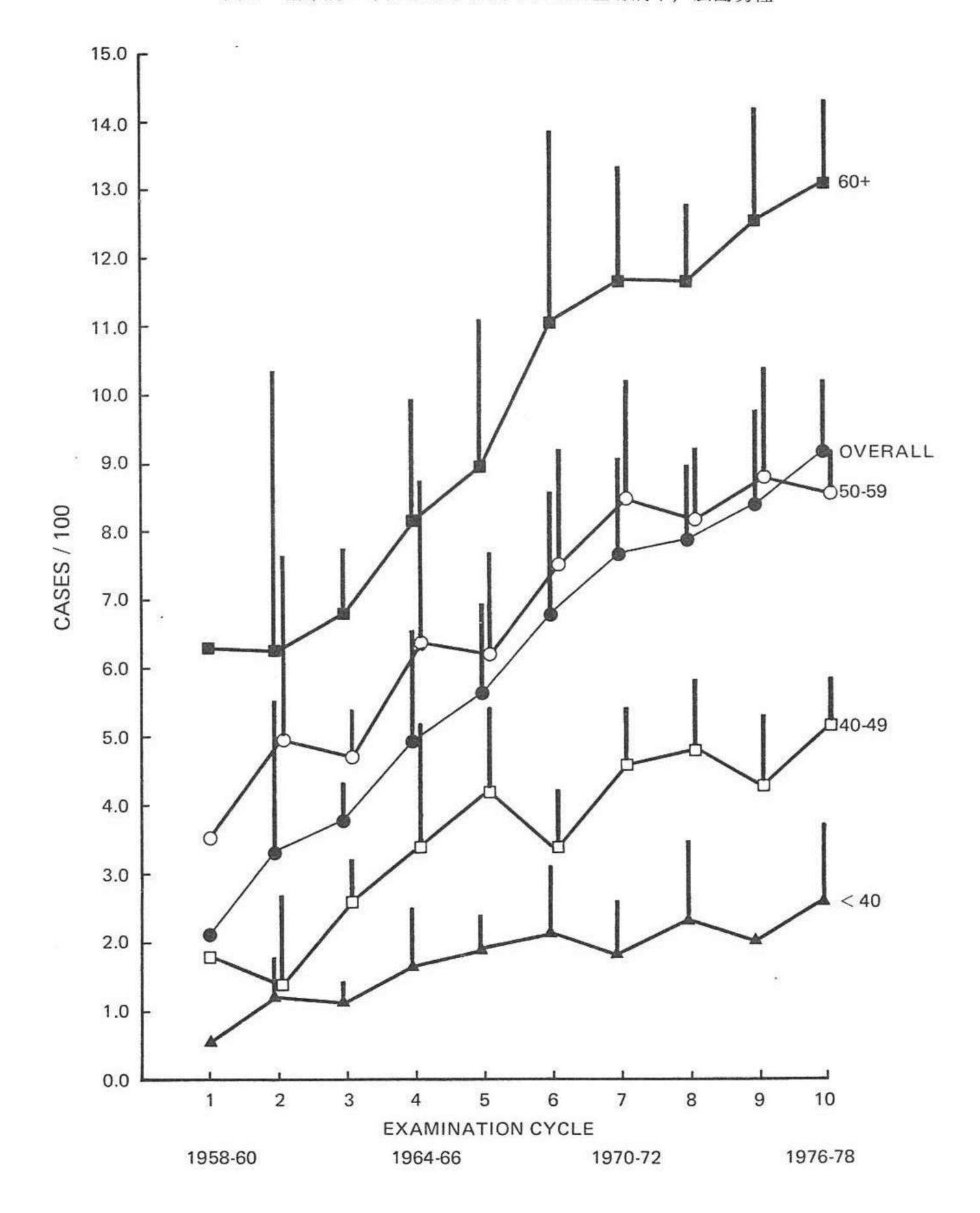
有病率の増加は,一定した年齢別発生率と集団の 老齢化によるという一つの解釈がある. NIDDM は 急激な致命的疾病ではないので, 生存する糖尿病 患者数は長期間には増加するかもしれない. これが 事実であれば, 集団が老齢化するにつれて有病率も 増加するであろう. しかし、他の解釈もあり得ること に留意すべきである.

広島及び長崎在住の日本人に関する20年間のコホート 研究で得た糖尿病,糖尿,及び蛋白尿に関する所見 を述べた. 驚いたことに、これまで日本人の糖尿病に 関する長期的な集団研究の報告はほとんどなかった. AHS のコホートに関する過去の短期研究2.3 では, 糖尿病及び糖尿の有病率は,長崎よりも広島の方が 高く,両市とも女性よりも男性の方が高いことが 認められた. また広島の対象者は、糖尿病の家族 歴がある場合, 性別に関係なく糖尿病の危険が増加 することも報告された.しかし、糖尿病患者における 蛋白尿率の増加については意見が分かれている.3,4 最近,NDDG10は以前のこれらの研究で用いられた ものより厳密な糖尿病の基準を提案した. したがって, 今回の研究は、NDDG の基準を用いた場合に過去の 所見がどのように変わるかを検討し、1978年までの AHS の糖尿病に関する所見を更新することを目的と した.

AHS の糖尿病に関連ある事実で注目に値するものが 二つある.一つは、欧米に比べ日本ではIDDMが 極めてまれなことである. 事実, 約20,000人からなる AHS の全コホート中に IDDM の症例はない. もう一つ は,第6周期(1968-70年)以降には,長崎ではほと

FIGURE 2 AGE-SPECIFIC OBSERVED AND ADJUSTED PREVALENCE RATES OF DIABETES, HIROSHIMA MALES

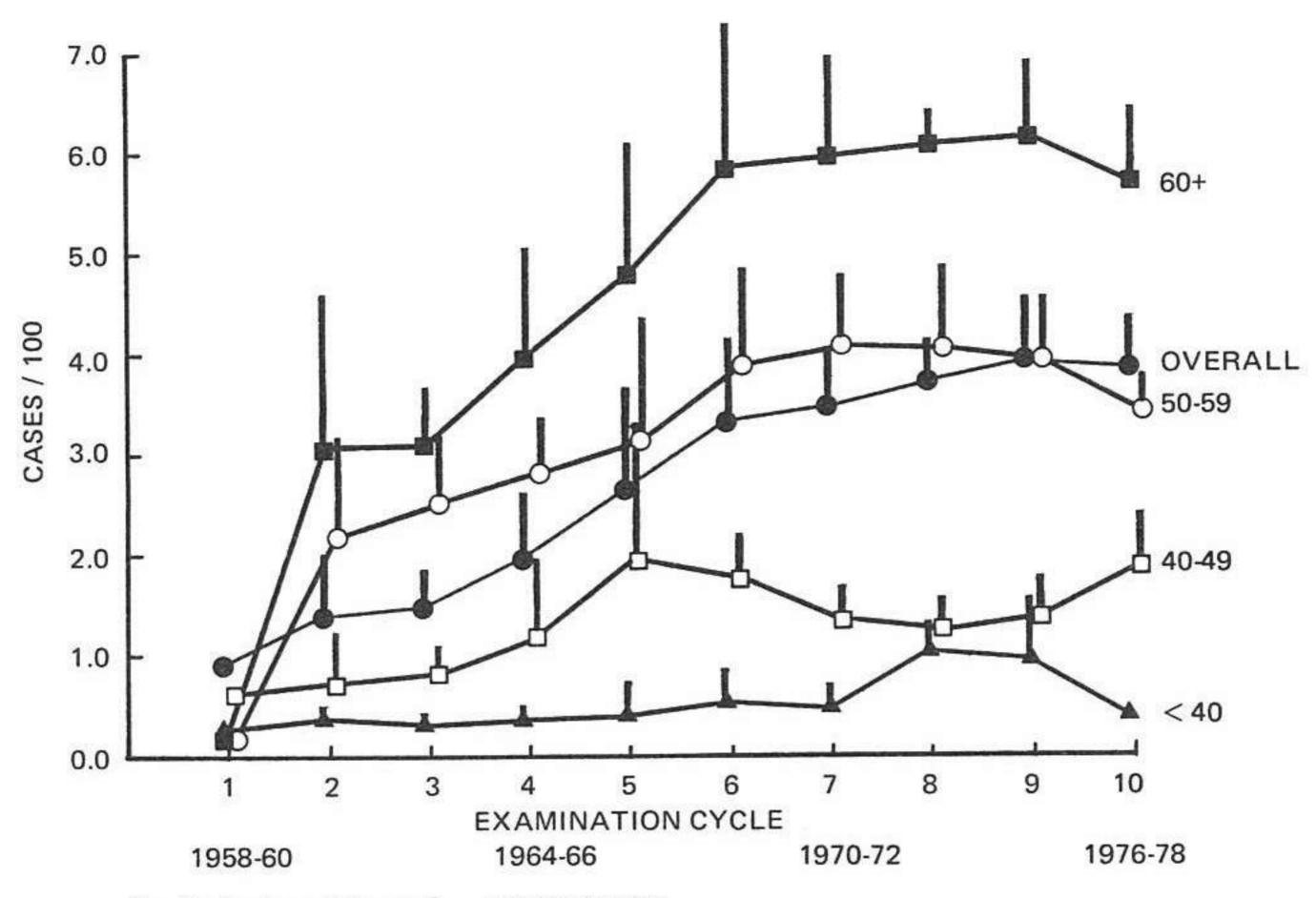
図2 糖尿病の年齢別観察有病率及び調整有病率,広島男性



Line segments connect observed prevalence rates for a given age-group. The top of a vertical bar rising from an observed prevalence rate indicates the corresponding adjusted prevalence rate. 線分は各年齢群の観察有病率を連結する、観察有病率から立ち上る縦棒の頂点は対応する調整有病率を示す.

FIGURE 3 AGE-SPECIFIC OBSERVED AND ADJUSTED PREVALENCE RATES OF DIABETES, HIROSHIMA FEMALES

図3 糖尿病の年齢別観察有病率及び調整有病率,広島女性



See footnotes of Figure 2. 図2の脚注参照.

of diabetes were restricted to the first 12 years (or six cycles) of the study, and Nagasaki-specific analyses were restricted to consideration of glycosuria, proteinuria, and diabetes family history rates.

The present findings on glycosuria generally agreed with those previously reported regarding this cohort. Over the 20 years of study, glycosuria was more prevalent in males than in females (p<.001 for both cities) and more prevalent in Hiroshima than in Nagasaki (p<.001 for both sexes). Among those subjects without known diabetes, glycosuria was more prevalent for both sexes (p<.001) in Hiroshima than in Nagasaki, but among diabetics there was little difference by city or sex. Thus, part of the observed differences by city and sex in diabetes morbidity may be attributed to the methodology used for detecting diabetes (namely, a routine glycosuria determination for primary screening). In fact, the observed difference of diabetes prevalence between the sexes might disappear or be reversed, if an OGTT were used as the primary (i.e., screening) test. 16 Another finding about glycosuria was that Hiroshima diabetics were glycosuric only about 46% of the time. This rate was unaffected by age, sex, て,糖尿病に関する両市間の比較は最初の12年間 (6周期)の研究に限られ,また長崎関係の解析は, 糖尿率,蛋白尿率,及び糖尿病家族歴の率の検討に 限定された.

糖尿に関する今回の所見は, このコホートについて 過去に報告されたものと大体一致する. 20年間の 研究を通じて、糖尿は女性よりも男性に多発し(両市 とも p<.001), また長崎よりも広島に多発した (男女 ともp<.001). 既知の糖尿病が認められない対象者 において, 糖尿は, 男女とも長崎より広島で有病率 が高かったが(p<.001), 糖尿病患者においては, 都市又は性別の差異はほとんどなかった.したがって, 糖尿病罹病率における都市及び性別の観察差異は, 糖尿病検出に用いた方法によるものかもしれない (すなわち、一次スクリーニングのための通常の 糖尿測定). 実際, 仮に OGTT を一次検査(つまり, スクリーニング)として用いたならば、男女間における 糖尿病有病率の観察差異は消失するか, 又は逆転 するかもしれない.16糖尿についてのもう一つの所見 は,広島の糖尿病患者で糖尿を示した者は,その 期間のわずか46%程度にすぎなかったということで ある.この率は年齢、性、及び肥満の有無によっては and presence or absence of obesity (Table 8). The main importance of this finding was twofold. It was used in estimating the size of the undiagnosed diabetic population and in explaining the success of routine urinalyses for diabetes prescreening.

Regarding proteinuria, the prevalence among diabetics was significantly higher than among all other subjects (p<.001 in Hiroshima, p≤.005 in Nagasaki). Among those subjects without known diabetes, the proteinuria rates were the same for all city-sex groups (p>.10, averaging 4.3%). For diabetics, within each city there was no differece in proteinuria rates between the sexes (p>.15). The diabetic proteinuria rates were 23.5% in Hiroshima and 18.7% in Nagasaki. These findings are in general agreement with those of Belsky et al.⁴

As would be expected, the main effect of using the NDDG diabetes criteria was to lower the prevalence rates in the AHS to about half those previously²⁻⁴ given. For cycle 1 (1958-60) Rudnick and Anderson² reported a prevalence of about 3% in Hiroshima in comparison to our 1.4%, whereas for cycle 5 (1966-68) Belsky et al⁴ reported a prevalence of about 6.3% compared to our 3.5%. For cycle 10 (1976-78), we find an overall prevalence of about 5.7%, which (when age and sex adjusted) closely agrees with recent prevalence estimates in Hiroshima agricultural workers.¹⁶

The age- and sex-specific incidence rates for diabetes were approximately constant in Hiroshima from cycle 4 (1964-66) onward. Unfortunately, in Nagasaki there were too few data for an analysis of incidence rates. The Hiroshima males had an overall incidence rate of about 1.2-1.3 cases/100 per cycle for cycles 4-10 (1964-78), while for the same period the Hiroshima females had an overall incidence (excluding cycle 5) of about 0.5-0.6 cases/100 per cycle. The glycosuria rate for the Hiroshima females was unusually high during cycle 5 (1966-68) and consequently so was the diabetes incidence rate.

The prevalence of diabetes in both sexes in Hiroshima was shown to increase over the course of study (p<.001). Unfortunately, the lack of adequate data precludes making any definite statements about temporal patterns in Nagasaki. The total prevalence in Hiroshima increased nearly 4-fold from 1.4 cases/100 in cycle 1 (1958-60) to 5.5 cases/100 in cycle 10 (1976-78). There were 705 individuals diagnosed as diabetic out of 13,718 Hiroshima sub-

影響されなかった(表8).この所見の主な重要性は 二つある.この所見は、診断が下されていない糖尿 病患者集団の規模の推定に用い、また、糖尿病の 前スクリーニングに、通常の尿検査が有効であること の説明に用いた.

蛋白尿については、糖尿病患者における有病率が、ほかのすべての対象者よりも有意に高かった (広島: p < .001, 長崎: p < .005). 既知の糖尿病の認められない対象者では、蛋白尿の率はすべての都市-性群において等しかった (p > .10, 平均 4.3%). 糖尿病患者に関しては、各都市内で蛋白尿率の男女差はなかった (p > .15). 糖尿病患者の蛋白尿率は、広島で23.5%、長崎で18.7%であった. これらの所見はBelsky 6^4 のものと大体一致している.

予想されたように、NDDG 糖尿病基準を使用すると、AHS の有病率が過去の例 $^{2-4}$ の約1/2 に低下した.第一周期 (1958-60年) については、Rudnick 及び Anderson 2 が、広島の有病率は我々の1.4%に対し、約3%であったと報告し、また第5 周期 (1966-68年) については、Belsky 6^4 が、我々の3.5%に対し 約6.3%の有病率を報告した.第10 周期 (1976-78年) については、全体で約5.7%の有病率が認められるが、これは、(年齢、性を調整した場合) 広島の農業 労働者 16 における最近の推定有病率にかなり一致している.

糖尿病の年齢及び性別発生率は、広島では第4周期 (1964-66年) 以降ほぼ一定していた。残念ながら、長崎では発生率の解析を行うにはデータがあまりにも少なかった。第4-10周期 (1964-78年) における広島の男性の全体の発生率は、1周期当たり約1.2-1.3 症例/100であり、同期間中の(第5周期を除く)広島の女性における全体の発生率は、1周期当たり約0.5-0.6症例/100であった。広島の女性の糖尿率は第5周期 (1966-68年) で異常に高く、その結果、糖尿病発生率もまた高かった。

広島の男女における糖尿病の有病率は、いずれも、研究期間を通して増加することが認められた(p<.001). 残念ながら、十分なデータがないために、長崎の経時的パターンについて明確な記述を行うことができない. 広島の総有病率は、第1周期(1958-60年)の1.4症例/100から第10周期(1976-78年)の5.5症例/100まで4倍近くも増加した. 少なくとも1回は検査を受けた広島の対象者13,718人のうち、

jects who were examined at least once. These increases are hypothesized to be the result of temporal changes in the age-sex distribution of the cohort combined with constant age- and sex-specific incidences. Over the 20 years of follow-up, the proportion of the cohort in the age-goups with high risk for diabetes increased. For example, in Hiroshima the percentage of persons aged 50 or more increased from 45% in males and 37% in females during cycle 1 to 64% in males and 73% in females during cycle 10. Using data from Bale and Entmacher, 17 it has been estimated 18 that in the State of Iowa the life expectancy of diabetics aged 40 or more is only about 10% less than that of similarly aged members of the general population. Since most of the AHS diabetics were aged 40 or more at diagnosis, the explanation that the increasing prevalences were primarily due to aging seems quite plausible.

For each cycle, an attempt was made to estimate the size of the undiagnosed diabetic population and to adjust our observed incidence and prevalence rates accordingly (Table 9). In terms of the number of cases missed, the same adjustments by cycle were used for the age- and sex-specific incidence and prevalence rates. Hence, the relative effect of the adjustments was much larger in the case of the incidence rates than for the prevalence rates. Because a routine urinalysis was the primary prescreening procedure for diabetes, it was shown in the Materials and Methods section that the observed incidence rates were too small by a factor about equal to the reciprocal of the probability of glycosuria in diabetics. For both sexes and all ages and weights, this adjustment factor was estimated to be 2.2. The ratio of observed prevalence to adjusted prevalence (for Hiroshima males) varied from 0.74 for cycle 4 to greater than 0.88 for cycle 10. If our assumption that the associated biases cancel each other out is correct, then the adjustments of the observed prevalences are probably fairly accurate. Consequently, the above values can be interpreted to mean that approximately 74% of the entire diabetic population had been diagnosed by the end of cycle 4 (1964-66) and that this percentage had increased to greater than 88% by the end of cycle 10 (1976-78). Thus, based on the present cohort study, it appears that continuous biennial routine urinalysis is an effective diabetes prescreening procedure.

There is a probabilistic reason for the effectiveness

糖尿病と診断された者が705人いた. これらの増加 は、コホートの年齢-性分布の経時的変化と、一定の 年齢・性別発生率とが結合した結果であると仮定 されている。20年間にわたる経過観察中、糖尿病の 危険率が高い年齢群のコホート内比率は増加した. 例えば、広島の50歳以上の人の割合は、第1周期 に男性45%, 女性37%であったものが, 第10周期 には男性64%, 女性73%まで増加した. Bale 及び Entmacher 17 のデータによって、Iowa 州の40歳以上 の糖尿病患者の寿命は,一般集団の同じ年齢群の それよりもわずか10%低いにすぎないことが推定され ている.18 AHS の糖尿病患者の大部分は, 診断時に 40歳以上であったので、有病率の増加は主として 加齢によるものであったとする説明は極めて妥当な ように思われる.

各周期ごとに、診断が下されていない糖尿病集団の 規模を推定し、それに従って我々の観察発生率と 有病率を調整する試みを行った(表9). 見落とした 症例数については,年齢・性別発生率並びに有病率に 対して, 周期別に同様の調整を用いた. したがって, 調整の相対的影響は有病率より発生率において, より大きかった. 通常の尿検査は、糖尿病を調べる 第一次前スクリーニング手続きであったので, 材料 及び方法の項に、糖尿病における糖尿の確率の逆数 にほぼ等しい分だけ観察発生率が小さくなりすぎた ことを示した. 男女とも, また全年齢及び体重に ついて、この調整係数は2.2と推定された. 観察有 病率の調整有病率に対する比(広島の男性の場合) は, 第4周期の0.74から第10周期の0.88以上まで 様々であった. 随伴している偏りが互いに相殺し合う という我々の仮定が正しいとすれば、 観察有病率 の調整は恐らくかなり正確であろう. したがって 上記の値は、全糖尿病患者集団の約74%が第4周期 (1964-66年)の終わりまでに診断されており、また この割合が、第10周期(1976-78年)の終わりまで には、88%以上に増加したことを意味するものと 解釈できる. このように、今回のコホート研究に よれば、2年ごとの通常尿検査の継続は、糖尿病の 前スクリーニングの方法として有効なものであると 思われる.

糖尿病の第一次スクリーニング手順の有効性について of the primary prescreening procedure for diabetes. は、確率論的理由がある。糖尿病患者が糖尿を示す

The probability of glycosuria in a diabetic was estimated to be about 0.46 without consideration of treatment. If this were the true value, the probability of nonglycosuria in an undiagnosed, asymptomatic diabetic would be 0.54. Because urinalyses were performed biennially and were considered statistically independent, the probability of an undiagnosed, asymptomatic diabetic having consecutive, negative urine specimens would be $(0.54)^n$. Consequently, the chances that an undiagnosed diabetic would be glycosuric at least once in two cycles would be 0.71 and in three cycles 0.84. Since most glycosurics without a prior diagnosis of diabetes were tested for diabetes, the prescreening procedure was quite effective, especially when looked at over the course of 20 years.

Because of the paucity of previous long-term population studies of diabetes among the Japanese, it seems appropriate to describe briefly some general characteristics of Japanese diabetics and to compare these characteristics with those of diabetics (NIDDM) in the West. Points of similarity include the average age at diagnosis (Table 4) and an increased relative risk of diabetes due to positive family history (Table 5). Also, even though glycosuria was more prevalent in males than in females, among glycosurics, diabetes was more prevalent in females than in males. In contrast to the West, the diabetics were predominantly male and not excessively overweight. Our observed 2:1 male predominance agrees with previous studies of the Japanese population^{1,14} and the AHS cohort.^{2,3} The difference in obesity levels between the AHS diabetics and western diabetics was quite striking. It has been reported¹⁰ that 60%-90% of those NIDDM in the West exceed 120% of IBW. In Hiroshima, however, it was observed that only 29% of the male diabetics and 44% of the female diabetics were in excess of 115% of IBW (Table 4).

It should be noted that these differences in obesity rates may be attributable to the much lower prevalence of obesity in the Japanese population than in the West.

The present study reports findings on diabetes, glycosuria, and proteinuria morbidity in the AHS. Since the AHS is a continuing study, future analyses of diabetes in this cohort, including possible complications of diabetes, are anticipated. 確率は、治療を考慮に入れないで約0.46であると 推定された.これば真の値であれば、診断の下され ていない無症状の糖尿病患者における非糖尿の確率 は0.54になる.尿検査は2年おきに実施され、統計 的に独立したものと考えられたので、診断の下され ていない無症状の糖尿病患者で尿検査の結果が連続 的に陰性であった者の確率は、(0.54) Pとなるで あろう.その結果、診断の下されていない糖尿病 患者が、2周期中に少なくとも1回糖尿を示す確率 は0.71となり、3周期中には0.84となる.以前に 糖尿病の診断が下されていない糖尿例の大部分に ついて糖尿病の検査を行ったので、前スクリーニング 手順は、特に20年の全期間を検討するときには極め て有効であった.

日本人の糖尿病に関する過去の長期集団研究は少な いので, 日本人の糖尿病患者の一般的特徴を簡単に 述べ,これらの特徴と欧米の糖尿病患者(NIDDM)の それを比較することは適切であると思われる. 類似点 は,診断時の平均年齢(表4)並びに陽性の家族歴 による糖尿病の相対危険度の増加(表5)などである. また、糖尿は女性よりも男性に多かったにもかかわ らず,糖尿を有する人たちの中では,糖尿病は男性 よりも女性に多く認められた. 欧米と対照的に, 糖尿病患者は男性に多く, 肥満例はさほど認められ なかった. 我々の観察した2対1の男性優勢比率は, 日本人集団1,14 及び AHS コホート2,3 に関する過去 の研究と一致する. AHS 糖尿病患者と欧米の糖尿病 患者との肥満度の差は、極めて顕著なものであった. 欧米では NIDDM の60%~90%が120% IBW を超え ていると報告されている.10しかし、広島では男性 糖尿病患者のわずか29%及び女性糖尿病患者の44% が 115% IBW を超過しているにすぎなかった (表4).

留意すべきことは、肥満率におけるこれらの差異は 日本人集団の肥満症の率が欧米に比べてはるかに低い ことによる可能性があるということである.

今回の研究では AHS における糖尿病,糖尿,及び蛋白尿の罹病率に関する所見を報告した. AHS は継続的な研究であるから,糖尿病の合併症も含め,このコホートの糖尿病に関する解析が将来も行われることを期待する.

APPENDIX TABLE 1 RATES (CASES/100) FOR GLYCOSURIA a AND PROTEINURIA b IN THE AHS AND AVERAGE AGES; BY AGE, CITY, AND SEX

付録表1 AHSにおける糖尿^a及び蛋白尿^bの率(症例/100)及び平均年齢;年齢,市,性別

| | · | | | 0 | C | ycle | | | | |
|-----------------|-------|------|---------|----------|---------|------|------|------|------|------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | | Hiroshin | na Male | | | | | |
| All ages | | | | | | | | | | |
| No. examined | 3495 | 3487 | 3267 | 3240 | 3038 | 2803 | 2584 | 2431 | 2166 | 1944 |
| Average age | 45.2 | 47.1 | 48.7 | 50.2 | 51.7 | 53.0 | 54.2 | 55.3 | 56.5 | 57.8 |
| Trace, 1+ glyco | - | 6.74 | 7.50 | 7.65 | 11.29 | 8.24 | 8.05 | 4.44 | 8.13 | 6.48 |
| 2+ - 4+ glyco | 9.10 | 2.49 | 2.60 | 2.75 | 3.69 | 3.39 | 3.02 | 1.97 | 3.69 | 3.65 |
| 1+, 2+ prot | - | - | - | 4.20 | 4.77 | 4.46 | 4.76 | 4.48 | 4.85 | 7.25 |
| 3+, 4+ prot | 7.10 | 6.54 | 5.60 | 0.52 | 0.53 | 0.61 | 0.58 | 0.70 | 0.55 | 0.67 |
| Age <30 | | | | | | | | | | |
| No. examined | 742 | 488 | 307 | 254 | 203 | 135 | 82 | 47 | 2 | 0 |
| Average age | 23.0 | 23.5 | 23.1 | 24.2 | 25.6 | 26.5 | 27.5 | 25.6 | 29.0 | - |
| Trace, 1+ glyco | - | 1.84 | 3.26 | 1.97 | 6.40 | 2.96 | 3.66 | 2.13 | 0.00 | - |
| 2+ - 4+ glyco | 4.18 | 1.02 | 0.10 | 0.79 | 1.48 | 0.00 | 0.00 | 0.00 | 0.00 | - |
| 1+, 2+ prot | - | _ | - | 0.79 | 2.96 | 2.22 | 2.44 | 0.00 | 0.00 | 4 |
| 3+, 4+ prot | 3.91 | 3.28 | 1.95 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | - |
| Age 30-39 | | | | | | | | | | |
| No. examined | 632 | 774 | 845 | 845 | 692 | 494 | 310 | 216 | 193 | 152 |
| Average age | 33.1 | 33.2 | 34.0 | 35.3 | 36.2 | 36.3 | 35.7 | 34.5 | 34.3 | 35.5 |
| Trace, 1+ glyco | (S#1) | 4.26 | 5.21 | 5.80 | 10.26 | 6.48 | 0.32 | 2.78 | 3.11 | 4.61 |
| 2+ - 4+ glyco | 6.80 | 1.29 | 1.42 | 1.78 | 2.02 | 0.40 | 4.52 | 1.39 | 1.04 | 0.66 |
| 1+, 2+ prot | - | - | - | 1.78 | 1.16 | 2.02 | 0.65 | 1.39 | 0.52 | 1.97 |
| 3+, 4+ prot | 3.16 | 2.71 | 1.66 | 0.00 | 0.00 | 0.00 | 0.32 | 0.46 | 0.00 | 0.00 |
| Age 40-49 | | | | | | | | | | |
| No. examined | 551 | 553 | 465 | 413 | 472 | 583 | 691 | 756 | 696 | 543 |
| Average age | 44.8 | 44.8 | 45.1 | 45.1 | 44.0 | 43.1 | 43.1 | 43.9 | 45.3 | 46.1 |
| Trace, 1+ glyco | - | 7.23 | 8.17 | 7.02 | 12.50 | 7.89 | 6.08 | 3.57 | 4.89 | 5.34 |
| 2+ - 4+ glyco | 8.89 | 1.27 | 3.01 | 3.15 | 2.33 | 4.12 | 3.76 | 2.51 | 3.30 | 2.76 |
| 1+, 2+ prot | - | - | - | 2.91 | 1.69 | 3.77 | 3.18 | 2.51 | 3.16 | 4.24 |
| 3+, 4+ prot | 3.81 | 3.25 | 2.80 | 0.00 | 0.42 | 0.00 | 0.00 | 0.26 | 0.29 | 0.37 |
| Age 50-59 | | | | | | | | | | |
| No. examined | 750 | 757 | 665 | 677 | 585 | 522 | 470 | 412 | 319 | 361 |
| Average age | 54.6 | 54.7 | 54.8 | 54.8 | 54.7 | 54.7 | 55.0 | 55.0 | 55.0 | 53.9 |
| Trace, 1+ glyco | - | 8.85 | 8.12 | 7.83 | 8.89 | 8.24 | 7.87 | 5.58 | 9.40 | 6.65 |
| 2+ - 4+ glyco | 11.20 | 3.30 | 2.71 | 3.69 | 5.47 | 3.83 | 3.40 | 1.70 | 3.13 | 3.88 |
| 1+, 2+ prot | | | H-SAMMA | 4.28 | 4.10 | 3.07 | 4.04 | 3.16 | 3.13 | 5.26 |
| 3+, 4+ prot | 7.60 | 7.00 | 5.86 | 0.30 | 0.51 | 0.38 | 0.43 | 0.24 | 0.00 | 0.28 |

APPENDIX TABLE 1 (Continued 続き)

| | Cycle | | | | | | | | | | |
|-----------------|-------|-------|-------|-----------|----------|-------|-------|------|-------|-------|--|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | |
| | | | | Hiroshim | na Male | | 5N | | | | |
| Age 60-69 | | | | | | | * | | | | |
| No. examined | 613 | 663 | 704 | 714 | 692 | 619 | 592 | 515 | 493 | 430 | |
| Average age | 63.6 | 63.8 | 64.1 | 64.4 | 64.3 | 64.2 | 64.6 | 64.6 | 64.5 | 64.5 | |
| Trace, 1+ glyco | - | 8.90 | 9.80 | 10.78 | 14.45 | 10.66 | 9.97 | 4.47 | 9.74 | 7.91 | |
| 2+ - 4+ glyco | 13.21 | 4.68 | 4.12 | 3.22 | 5.64 | 5.17 | 3.89 | 1.75 | 5.68 | 3.02 | |
| 1+, 2+ prot | _ | - | 72 | 6.16 | 8.53 | 5.65 | 5.57 | 6.41 | 6.09 | 6.05 | |
| 3+, 4+ prot | 12.56 | 11.61 | 9.09 | 1.40 | 0.72 | 0.65 | 0.68 | 0.78 | 0.61 | 0.93 | |
| Age 70+ | | | | | | | | | | | |
| No. examined | 207 | 252 | 281 | 337 | 394 | 450 | 439 | 485 | 463 | 458 | |
| Average age | 73.9 | 74.4 | 74.2 | 74.5 | 74.6 | 74.6 | 74.8 | 75.2 | 75.4 | 75.7 | |
| Trace, 1+ glyco | - | 10.71 | 10.68 | 10.39 | 12.18 | 8.89 | 12.07 | 5.77 | 12.53 | 6.99 | |
| 2+ - 4+ glyco | 14.49 | 3.57 | 4.27 | 3.26 | 3.30 | 3.78 | 2.05 | 2.06 | 3.67 | 6.11 | |
| 1+, 2+ prot | - | 8= | - | 10.09 | 10.15 | 8.67 | 10.25 | 8.45 | 9.07 | 15.28 | |
| 3+, 4+ prot | 21.26 | 17.06 | 16.73 | 1.48 | 1.52 | 2.44 | 1.82 | 1.86 | 1.51 | 1.31 | |
| | | | | (T:L: | . Ela | | | | | | |
| All ages | | | 1 | HIIOSHIII | a Female | | | | | | |
| No. examined | 6052 | 6199 | 5882 | 5871 | 5644 | 5384 | 4914 | 4708 | 4312 | 377 | |
| Average age | 43.8 | 45.7 | 47.6 | 49.2 | 50.9 | 52.5 | 54.2 | 55.4 | 56.6 | 57. | |
| Trace, 1+ glyco | - | 1.40 | 1.62 | 1.92 | 2.96 | 2.34 | 1.91 | 1.66 | 2.02 | 1.9 | |
| 2+ - 4+ glyco | 4.11 | 1.03 | 0.87 | 0.99 | 1.38 | 1.73 | 1.42 | 0.91 | 1.69 | 1.5 | |
| 1+, 2+ prot | _ | 12 | 92 | 3.78 | 3.58 | 3.58 | 3.95 | 4.06 | 4.24 | 5.8 | |
| 3+, 4+ prot | 6.03 | 5.81 | 4.96 | 0.65 | 0.66 | 0.45 | 0.47 | 0.49 | 0.56 | 0.6 | |
| Age <30 | | | | | | | | | | | |
| No. examined | 1079 | 752 | 473 | 384 | 305 | 212 | 130 | 56 | 3 | | |
| Average age | 23.5 | 23.9 | 23.5 | 24.3 | 25.5 | 26.8 | 27.7 | 28.5 | 29.0 | | |
| Trace, 1+ glyco | - | 1.06 | 0.63 | 2.08 | 1.31 | 1.42 | 0.00 | 0.00 | 0.00 | 100.0 | |
| 2+ - 4+ glyco | 4.26 | 0.13 | 0.42 | 0.26 | 0.33 | 0.00 | 0.77 | 0.00 | 0.00 | 0.0 | |
| 1+, 2+ prot | - | 82 | - | 1.56 | 2.30 | 3.30 | 3.85 | 1.79 | 0.00 | 0.0 | |
| 3+, 4+ prot | 3.34 | 3.86 | 3.59 | 0.26 | 0.33 | 0.00 | 0.77 | 0.00 | 0.00 | 0.0 | |
| Age 30-39 | | | | | | | | | | | |
| No. examined | 1572 | 1719 | 1594 | 1365 | 1019 | 739 | 474 | 306 | 297 | 22 | |
| Average age | 34.4 | 34.4 | 34.8 | 35.6 | 36.0 | 36.2 | 35.9 | 34.5 | 34.5 | 35. | |
| Trace, 1+ glyco | - | 1.16 | 1.00 | 1.10 | 2.45 | 2.03 | 1.27 | 1.31 | 1.01 | 1.3 | |
| 2+ - 4+ glyco | 3.94 | 0.35 | 0.50 | 0.22 | 0.59 | 0.54 | 1.05 | 0.33 | 0.34 | 0.4 | |
| 1+, 2+ prot | - | - | - | 2.12 | 0.98 | 1.35 | 1.69 | 1.31 | 1.68 | 1.7 | |
| 3+, 4+ prot | 2.80 | 2.73 | 2.63 | 0.07 | 0.49 | 0.41 | 0.63 | 0.33 | 0.00 | 0.8 | |

APPENDIX TABLE 1 (Continued 続き)

| <u>2-</u> | | | | | Су | cle | - | | | |
|-----------------------|-------|-------|--------------------|-----------|---------|------|------|------|------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | Н | Iiroshima | Female | | | | | |
| Age Age 40-49 | | | | | | | | | | |
| No. examined | 1134 | 1163 | 1154 | 1313 | 1464 | 1545 | 1480 | 1483 | 1147 | 799 |
| Average age | 44.7 | 44.5 | 44.2 | 43.9 | 44.0 | 44.2 | 44.5 | 44.9 | 45.5 | 46.1 |
| Trace, 1+ glyco | - | 1.03 | 1.30 | 1.45 | 2.53 | 2.01 | 1.49 | 1.15 | 0.96 | 1.25 |
| 2+ - 4+ glyco | 3.09 | 0.95 | 0.17 | 0.61 | 1.09 | 1.42 | 0.88 | 0.61 | 1.22 | 1.00 |
| 1+, 2+ prot | - | - | | 1.83 | 1.98 | 1.75 | 2.36 | 2.97 | 2.44 | 3.13 |
| 3+, 4+ prot | 4.76 | 3.35 | 2.17 | 0.38 | 0.34 | 0.32 | 0.34 | 0.34 | 0.78 | 0.50 |
| Age 50-59 | | | | | | | | | | |
| No. examined | 1280 | 1349 | 1260 | 1222 | 1134 | 1079 | 1027 | 1050 | 1130 | 1154 |
| Average age | 54.5 | 54.6 | 54.6 | 54.7 | 54.8 | 54.6 | 54.5 | 54.2 | 53.8 | 53.9 |
| Trace, 1+ glyco | - | 1.41 | 1.75 | 1.88 | 2.47 | 2.22 | 1.56 | 1.43 | 2.12 | 1.82 |
| 2+ - 4+ glyco | 4.38 | 1.48 | 1.83 | 1.15 | 1.50 | 1.76 | 1.36 | 0.67 | 1.86 | 1.21 |
| 1+, 2+ prot | :- | - | (-1 0) | 3.68 | 3.26 | 3.52 | 2.63 | 2.57 | 3.01 | 4.42 |
| 3+, 4+ prot | 7.11 | 6.67 | 5.08 | 0.65 | 0.35 | 0.46 | 0.10 | 0.38 | 0.35 | 0.35 |
| Age 60-69 | | | | | | | | | | |
| No. examined | 710 | 899 | 1034 | 1147 | 1197 | 1180 | 1061 | 1012 | 931 | 834 |
| Average age | 63.6 | 63.4 | 63.8 | 64.1 | 64.3 | 64.4 | 64.5 | 64.5 | 64.5 | 64.6 |
| Trace, 1+ glyco | = | 1.89 | 2.22 | 3.14 | 3.26 | 3.47 | 2.54 | 1.48 | 2.58 | 2.64 |
| 2+ - 4+ glyco | 5.35 | 1.89 | 1.16 | 2.27 | 2.51 | 2.63 | 2.26 | 1.78 | 2.26 | 2.28 |
| 1+, 2+ prot | - | - | - | 5.93 | 5.10 | 5.00 | 6.03 | 4.64 | 4.73 | 6.00 |
| 3+, 4+ prot | 9.44 | 9.01 | 7.93 | 1.13 | 1.34 | 0.59 | 0.38 | 0.49 | 0.54 | 0.72 |
| Age 70+ | | | | | | | | | | |
| No. examined | 270 | 317 | 367 | 440 | 525 | 629 | 742 | 801 | 804 | 763 |
| Average age | 74.4 | 75.4 | 75.1 | 75.0 | 74.8 | 74.8 | 74.8 | 75.1 | 75.3 | 75.8 |
| Trace, 1+ glyco | - | 3.47 | 4.36 | 2.73 | 6.48 | 1.91 | 3.10 | 3.37 | 3.11 | 2.49 |
| 2+ - 4+ glyco | 4.33 | 2.84 | 1.09 | 1.36 | 1.52 | 2.70 | 1.75 | 1.00 | 1.99 | 2.69 |
| 1+, 2+ prot | | 27 | - 1 | 11.36 | 11.05 | 8.27 | 7.41 | 8.49 | 8.96 | 11.66 |
| 3+, 4+ prot | 26.35 | 23.34 | 16.89 | 2.27 | 1.14 | 0.64 | 1.21 | 1.00 | 0.75 | 1.05 |
| | | | | | | | | | | |
| All ages | | | | Nagasal | ki Male | | | | | |
| All ages No. examined | 1242 | 1730 | 1763 | 1653 | 1549 | 1362 | 1392 | 1292 | 1274 | 116 |
| | 39.6 | 42.2 | 44.1 | 46.0 | 47.4 | 49.0 | 50.2 | 52.1 | 53.5 | 55. |
| Average age | 37.0 | 2.72 | 4.65 | 3.63 | 6.91 | 7.27 | 3.52 | 2.32 | 2.67 | 3.9 |
| Trace, 1+ glyco | 3.46 | 1.56 | 1.93 | 1.75 | 2.52 | 3.67 | 2.66 | 1.78 | 2.43 | 3.1 |
| 2+ - 4+ glyco | | | | 3.15 | 3.10 | 3.96 | 4.60 | 5.42 | 6.75 | 5.6 |
| 1+, 2+ prot | 175 | 2 77 | 3.12 | 0.00 | 0.26 | 0.37 | 0.14 | 0.23 | 0.78 | 0.5 |
| 3+, 4+ prot | 4.75 | 2.77 | 3.12 | 0.00 | 0.20 | 0.57 | 0.14 | 0.23 | 0.70 | 0.5 |

APPENDIX TABLE 1 (Continued 続き)

| | Cycle | | | | | | | | | | |
|-----------------|-----------|-------|-------|----------|-------|--------------|-------|--------|-------|-----|--|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | |
| | | | | Nagasaki | Male | | | | | | |
| Age <30 | 206 | 272 | 257 | 100 | 160 | 02 | 55 | 20 | 0 | , | |
| No. examined | 396 | 372 | 257 | 189 | 168 | 92 | 27.9 | 28.8 | 0 | (| |
| Average age | 22.5 | 24.1 | 23.6 | 24.7 | 26.0 | 27.0 4.35 | 1.82 | 5.00 | - | | |
| Trace, 1+ glyco | 1.50 | 0.54 | 3.50 | 1.06 | 2.38 | | | 5.00 | - | | |
| 2+ - 4+ glyco | 1.52 | 0.27 | 0.00 | 0.00 | 1.19 | 2.17 | 1.82 | 5.00 | = | | |
| 1+, 2+ prot | | - 15 | 1 50 | 1.06 | 1.79 | 3.26 | 1.82 | | 5 | | |
| 3+, 4+ prot | 2.53 | 2.15 | 1.56 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 9 | | |
| Age 30-39 | | | | | | | | 1912/9 | | | |
| No. examined | 238 | 461 | 549 | 527 | 448 | 329 | 269 | 181 | 168 | 13 | |
| Average age | 32.9 | 32.8 | 33.8 | 34.9 | 35.9 | 36.1 | 35.4 | 34.4 | 34.7 | 36. | |
| Trace, 1+ glyco | - | 3.25 | 4.37 | 4.17 | 7.14 | 4.56 | 2.97 | 1.10 | 0.60 | 2.9 | |
| 2+ - 4+ glyco | 1.68 | 0.87 | 1.82 | 0.38 | 2.23 | 3.04 | 1.86 | 1.10 | 1.10 | 1.4 | |
| 1+, 2+ prot | - | - | - | 1.52 | 2.01 | 3.04 | 2.23 | 1.10 | 2.38 | 2.9 | |
| 3+, 4+ prot | 2.94 | 1.30 | 1.64 | 0.00 | 0.22 | 0.00 | 0.00 | 0.00 | 0.00 | 0.7 | |
| Age 40-49 | | | | | | | | | | | |
| No. examined | 223 | 298 | 282 | 248 | 245 | 281 | 399 | 450 | 452 | 37 | |
| Average age | 44.5 | 45.0 | 45.3 | 45.3 | 44.3 | 42.7 | 42.9 | 43.8 | 44.9 | 45 | |
| Trace, 1+ glyco | - | 3.36 | 4.61 | 4.84 | 5.31 | 8.90 | 3.51 | 4.22 | 3.10 | 5.1 | |
| 2+ - 4+ glyco | 4.04 | 2.35 | 3.19 | 3.63 | 2.04 | 2.85 | 2.76 | 2.00 | 3.32 | 2.9 | |
| 1+, 2+ prot | - | 2 | -1 | 2.02 | 1.63 | 1.42 | 3.26 | 3.33 | 5.53 | 5.4 | |
| 3+, 4+ prot | 2.69 | 2.68 | 2.13 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.22 | 0.2 | |
| Age 50-59 | | | | | | | | | | | |
| No. examined | 236 | 331 | 343 | 323 | 306 | 299 | 281 | 222 | 213 | 21 | |
| Average age | 54.6 | 54.7 | 54.5 | 54.4 | 54.5 | 54.2 | 55.1 | 55.4 | 55.3 | 54 | |
| Trace, 1+ glyco | - | 2.72 | 4.66 | 2.79 | 6.21 | 8.03 | 1.78 | 1.80 | 0.94 | 3.2 | |
| 2+ - 4+ glyco | 5.51 | 2.42 | 1.17 | 2.79 | 3.27 | 5.69 | 2.49 | 1.35 | 1.41 | 1.8 | |
| 1+, 2+ prot | | - | - | 3.10 | 2.61 | 4.35 | 2.85 | 4.05 | 3.76 | 1.3 | |
| 3+, 4+ prot | 7.20 | 3.32 | 3.21 | 0.00 | 0.00 | 0.33 | 0.36 | 0.45 | 0.47 | 0.0 | |
| Age 60-69 | | | | | | | | | | | |
| No. examined | 133 | 228 | 276 | 284 | 292 | 263 | 258 | 248 | 259 | 2 | |
| Average age | 63.7 | 63.5 | 63.7 | 63.9 | 64.3 | 64.3 | 64.6 | 64.3 | 64.4 | 64 | |
| Trace, 1+ glyco | - | 3.95 | 6.16 | 3.52 | 8.90 | 6.84 | 5.43 | 1.61 | 3.09 | 3.6 | |
| 2+ - 4+ glyco | 6.77 | 3.07 | 3.99 | 3.17 | 3.42 | 3.80 | 2.71 | 1.61 | 2.32 | 3.2 | |
| 1+, 2+ prot | 6200m3 82 | F-2 | - | 6.69 | 5.48 | 4.94 | 7.75 | 8.47 | 8.11 | 6. | |
| 3+, 4+ prot | 11.28 | 3.95 | 6.52 | 0.00 | 0.34 | 1.14 | 0.39 | 0.40 | 1.54 | 0. | |
| Age 70+ | | | | | | | | | | | |
| No. examined | 16 | 40 | 56 | 82 | 90 | 98 | 130 | 171 | 182 | 1 | |
| Average age | 72.9 | 73.9 | 72.9 | 73.3 | 73.6 | 73.4 | 73.7 | 73.9 | 74.6 | 75 | |
| Trace, 1+ glyco | - | 5.0 | 5.36 | 6.10 | 14.44 | 13.27 | 5.38 | 0.00 | 4.95 | 3. | |
| 2+ - 4+ glyco | 12.5 | 5.0 | 0.00 | 0.00 | 2.22 | 3.06 | 4.62 | 2.34 | 2.75 | 5. | |
| 1+, 2+ prot | 12.5 | | - | 9.76 | 8.89 | 11.22 | 12.31 | 12.87 | 15.38 | 11. | |
| 3+, 4+ prot | 25.00 | 15.00 | 12.50 | 0.00 | 2.22 | 1.02 | 0.00 | 0.58 | 2.20 | 2. | |

APPENDIX TABLE 1 (Continued 続き)

| | Cycle | | | | | | | | | | |
|-----------------|-------|------|------|----------|--------|------|------|------|-------|------|--|
| • | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | |
| | | 2 | 1 | Vagasaki | Female | | | | | | |
| All ages | | | | | | | | | | | |
| No. examined | 1716 | 2443 | 2480 | 2392 | 2285 | 2030 | 2136 | 2000 | 1966 | 1805 | |
| Average age | 37.5 | 39.4 | 41.4 | 43.1 | 44.8 | 46.4 | 48.5 | 50.1 | 51.7 | 53.2 | |
| Trace, 1+ glyco | | 0.49 | 0.93 | 1.46 | 1.44 | 3.00 | 0.70 | 0.70 | 0.81 | 1.33 | |
| 2+ - 4+ glyco | 1.69 | 0.53 | 0.65 | 0.84 | 1.09 | 1.38 | 0.94 | 0.50 | 1.37 | 0.94 | |
| 1+, 2+ prot | - | - | - | 2.88 | 3.50 | 3.74 | 3.46 | 3.80 | 4.78 | 4.49 | |
| 3+, 4+ prot | 5.71 | 2.21 | 2.38 | 0.21 | 0.13 | 0.10 | 0.28 | 0.30 | 0.61 | 0.55 | |
| Age <30 | | | | | | | | | | | |
| No. examined | 517 | 451 | 349 | 261 | 176 | 115 | 70 | 35 | 3 | C | |
| Average age | 23.1 | 24.4 | 24.3 | 25.0 | 25.8 | 26.6 | 27.4 | 28.5 | 29.0 | 9 | |
| Trace, 1+ glyco | - | 0.67 | 0.29 | 1.15 | 1.14 | 0.87 | 0.00 | 0.00 | 0.00 | 120 | |
| 2+ - 4+ glyco | 2.71 | 0.44 | 0.57 | 0.77 | 0.57 | 1.74 | 0.00 | 0.00 | 0.00 | | |
| 1+, 2+ prot | - | - | - | 2.68 | 2.27 | 2.61 | 0.00 | 0.00 | 0.00 | | |
| 3+, 4+ prot | 2.71 | 1.33 | 1.72 | 0.38 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | | |
| Age 30-39 | | | | | | | | | | | |
| No. examined | 566 | 1071 | 1032 | 902 | 752 | 519 | 351 | 248 | 209 | 144 | |
| Average age | 33.8 | 33.9 | 34.7 | 35.3 | 35.8 | 36.1 | 35.7 | 35.6 | 35.2 | 35.9 | |
| Trace, 1+ glyco | - | 0.09 | 0.78 | 0.78 | 1.33 | 2.70 | 0.85 | 0.81 | 0.96 | 2.08 | |
| 2+ - 4+ glyco | 1.59 | 0.00 | 0.29 | 0.22 | 0.27 | 0.58 | 0.00 | 0.40 | 1.91 | 0.00 | |
| 1+, 2+ prot | - | - | - | 1.55 | 2.13 | 2.70 | 1.99 | 2.82 | 2.39 | 2.78 | |
| 3+, 4+ prot | 4.06 | 1.31 | 1.26 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.48 | 0.00 | |
| Age 40-49 | | | | | | | | | | | |
| No. examined | 271 | 381 | 488 | 582 | 701 | 763 | 929 | 898 | 800 | 622 | |
| Average age | 44.1 | 43.7 | 43.7 | 43.6 | 43.6 | 43.7 | 43.8 | 44.7 | 45.3 | 46.0 | |
| Trace, 1+ glyco | - | 0.79 | 0.61 | 1.72 | 1.71 | 3.41 | 0.65 | 0.89 | 0.25 | 0.80 | |
| 2+ - 4+ glyco | 1.11 | 0.52 | 1.02 | 0.69 | 1.14 | 1.70 | 0.65 | 0.45 | 0.63 | 0.64 | |
| 1+, 2+ prot | _ | - | - | 2.75 | 2.28 | 2.10 | 1.95 | 2.34 | 3.38 | 0.16 | |
| 3+, 4+ prot | 4.80 | 2.89 | 2.66 | 0.17 | 0.29 | 0.00 | 0.00 | 0.00 | 0.13 | 0.80 | |
| Age 50-59 | | | | | | | | | | | |
| No. examined | 215 | 309 | 308 | 308 | 303 | 282 | 372 | 406 | 513 | 602 | |
| Average age | 54.3 | 54.5 | 54.6 | 54.3 | 54.4 | 54.2 | 53.6 | 53.8 | 53.6 | 53.7 | |
| Trace, 1+ glyco | ≌ | 0.97 | 0.65 | 1.30 | 0.99 | 1.77 | 0.54 | 0.25 | 0.97 | 0.83 | |
| 2+ - 4+ glyco | 0.93 | 1.29 | 0.32 | 1.62 | 2.64 | 1.77 | 2.15 | 0.25 | 1.36 | 1.33 | |
| 1+, 2+ prot | | # | Ħ | 4.22 | 3.63 | 5.32 | 3.76 | 3.94 | 2.34 | 4.15 | |
| 3+, 4+ prot | 11.16 | 4.85 | 4.22 | 0.32 | 0.00 | 0.71 | 0.81 | 0.74 | 1.17 | 0.66 | |
| Age 60-69 | | | | | | | | | | | |
| No. examined | 115 | 168 | 219 | 245 | 235 | 232 | 256 | 244 | 247 | 253 | |
| Average age | 63.9 | 64.0 | 63.8 | 64.0 | 63.9 | 64.3 | 64.4 | 64.5 | 64.4 | 64.2 | |
| Trace, 1+ glyco | - | 0.60 | 2.74 | 2.86 | 2.13 | 3.88 | 1.56 | 1.23 | 0.40 | 2.37 | |
| 2+ - 4+ glyco | 0.87 | 2.38 | 1.37 | 2.04 | 1.28 | 1.72 | 1.56 | 0.82 | 2.43 | 0.79 | |
| 1+, 2+ prot | - | - | - | 3.67 | 6.38 | 4.74 | 6.25 | 6.15 | 10.53 | 7.11 | |
| 3+, 4+ prot | 10.43 | 1.79 | 3.20 | 0.82 | 0.43 | 0.00 | 0.39 | 0.41 | 0.40 | 1.58 | |

APPENDIX TABLE 1 (Continued 続き)

| | Cycle | | | | | | | | | | | |
|-----------------|-------|------|------|----------|--------|-------|-------|-------|-------|-------|--|--|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | | |
| | | | 1 | Nagasaki | Female | | | | | | | |
| Age 70+ | | | | | | | | | | | | |
| No. examined | 32 | 63 | 84 | 94 | 118 | 119 | 163 | 169 | 194 | 184 | | |
| Average age | 73.3 | 73.7 | 74.1 | 73.9 | 73.7 | 74.4 | 74.8 | 74.7 | 74.9 | 74.8 | | |
| Trace, 1+ glyco | - | 1.59 | 3.57 | 4.26 | 0.85 | 5.04 | 0.00 | 0.00 | 3.09 | 2.72 | | |
| 2+, 4+ glyco | 0.00 | 1.59 | 2.38 | 2.13 | 2.54 | 0.84 | 1.23 | 1.18 | 2.58 | 1.63 | | |
| 1+, 2+ prot | - | - | = | 10.64 | 15.25 | 14.29 | 11.66 | 10.06 | 12.37 | 11.41 | | |
| 3+, 4+ prot | 37.5 | 7.94 | 8.33 | 0.00 | 0.00 | 0.00 | 1.23 | 1.18 | 1.55 | 0.54 | | |

a For cycle 1, glycosuria was recorded as trace or greater.

APPENDIX TABLE 2 OBSERVED AND ADJUSTED MORBIDITY RATES (CASES/100) FOR DIABETES BY AGE, CITY, AND SEX

付録表 2 糖尿病の観察及び調整罹病率 (症例/100): 年齢, 都市, 及び性別

| | | | | | C | ycle | | | | |
|---------------------|-------------|------|-------|---------|------|------|------|------|------|------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | | Hiros | shima N | [ale | | | | | |
| Age <40 | | | | | | | | | | |
| Observed Incidence | - | 1.0 | .2 | .5 | .3 | .6 | .5 | .8 | .0 | .7 |
| Adjusted Incidence | - | 2.5 | .4 | 1.2 | .8 | 1.7 | 1.3 | 2.0 | .0 | 1.8 |
| Observed Prevalence | .5 | 1.2 | 1.1 | 1.7 | 1.9 | 2.1 | 1.8 | 2.3 | 2.0 | 2.6 |
| Adjusted Prevalence | - | 2.7 | 1.4 | 2.5 | 2.4 | 3.1 | 2.6 | 3.5 | 2.0 | 3.7 |
| Age 40-49 | | | | | | | | | | |
| Observed Incidence | - | .4 | .6 | 2.0 | 1.3 | .9 | .9 | 1.1 | 1.0 | .8 |
| Adjusted Incidence | - | 1.7 | 1.3 | 3.8 | 2.6 | 1.7 | 1.8 | 2.1 | 2.0 | 1.5 |
| Observed Prevalence | 1.8 | 1.4 | 2.6 | 3.4 | 4.2 | 3.4 | 4.6 | 4.8 | 4.3 | 5.2 |
| Adjusted Prevalence | () | 1.8 | 3.2 | 5.2 | 5.4 | 4.2 | 5.4 | 5.8 | 5.3 | 5.8 |
| Age 50-59 | | | | | | | | | | |
| Observed Incidence | - | 2.7 | .8 | 2.5 | 1.6 | 1.8 | 1.8 | 1.0 | 1.7 | .6 |
| Adjusted Incidence | - | 5.5 | 1.6 | 5.0 | 3.3 | 3.7 | 3.7 | 2.1 | 3.4 | 1.2 |
| Observed Prevalence | 3.5 | 5.0 | 4.7 | 6.4 | 6.2 | 7.5 | 8.5 | 8.2 | 8.8 | 8.6 |
| Adjusted Prevalence | - | 7.7 | 5.4 | 8.8 | 7.7 | 9.2 | 10.2 | 9.2 | 10.4 | 9.2 |
| Age 60+ | | | | | | | | | | |
| Observed Incidence | - | 3.4 | .8 | 1.6 | 1.8 | 2.5 | 1.5 | 1.0 | 1.5 | 1.2 |
| Adjusted Incidence | - | 7.5 | 1.7 | 3.6 | 4.0 | 5.5 | 3.4 | 2.2 | 3.4 | 2.6 |
| Observed Prevalence | 6.3 | 6.3 | 6.8 | 8.2 | 9.0 | 11.1 | 11.7 | 11.7 | 12.6 | 13.1 |
| Adjusted Prevalence | | 10.4 | 7.8 | 10.0 | 11.1 | 13.9 | 13.4 | 12.8 | 14.2 | 14.3 |

第1周期については、糖尿は極微量又はそれ以上として記録されている.

b For cycles 1-3, proteinuria was recorded as trace or greater.

第1-3周期については、蛋白尿は極微量又はそれ以上として記録されている。

APPENDIX TABLE 2 (Continued 続き)

| | | | | | Су | cle | | | | |
|---------------------|-----|-----|---------|---------|-----|-----|-----|-----|-----|-----|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| | | Hir | oshim | a Fema | ale | | | | | |
| Age <40 | | | | | | | | | | |
| Observed Incidence | 140 | .1 | .1 | .1 | .3 | .2 | .2 | .3 | .3 | .(|
| Adjusted Incidence | 120 | .3 | .1 | .3 | .7 | .5 | .4 | .6 | .8 |). |
| Observed Prevalence | .3 | .4 | .3 | .4 | .4 | .6 | .5 | 1.1 | 1.0 | .4 |
| Adjusted Prevalence | - | .5 | .4 | .5 | .8 | .9 | .7 | 1.4 | 1.4 | -4 |
| Age 40-49 | | | | | | | | | | |
| Observed Incidence | - | .4 | .3 | .6 | 1.0 | .3 | .2 | .2 | .3 | |
| Adjusted Incidence | - | 1.0 | .6 | 1.4 | 2.3 | .6 | .5 | .5 | .6 | |
| Observed Prevalence | .6 | .7 | .8 | 1.2 | 2.0 | 1.8 | 1.4 | 1.3 | 1.4 | 1.9 |
| Adjusted Prevalence | - | 1.3 | 1.1 | 2.0 | 3.3 | 2.2 | 1.7 | 1.6 | 1.8 | 2. |
| Age 50-60 | | | | | | | | | | |
| Observed Incidence | - | .8 | .6 | .5 | 1.1 | .9 | .6 | .7 | .6 | |
| Adjusted Incidence | - | 1.8 | 1.2 | 1.1 | 2.4 | 1.9 | 1.3 | 1.5 | 1.2 | |
| Observed Prevalence | .2 | 2.2 | 2.5 | 2.8 | 3.1 | 3.9 | 4.1 | 4.1 | 4.0 | 3. |
| Adjusted Prevalence | - | 3.2 | 3.2 | 3.4 | 4.4 | 4.9 | 4.8 | 4.9 | 4.6 | 3. |
| Age 60+ | | | | | | | | | | |
| Observed Incidence | - | 1.4 | .6 | 1.0 | 1.2 | 1.3 | .9 | .4 | .8 | |
| Adjusted Incidence | - | 3.0 | 1.2 | 2.2 | 2.5 | 2.8 | 1.9 | .7 | 1.8 | 1. |
| Observed Prevalence | .20 | 3.1 | 3.1 | 4.0 | 4.8 | 5.9 | 6.0 | 6.1 | 6.2 | 5. |
| Adjusted Prevalence | - | 4.6 | 3.7 | 5.1 | 6.1 | 7.3 | 7.0 | 6.4 | 7.0 | 6. |
| | | N | lagasal | ci Male | e | | | | | |
| Age <40 | | | 575 | | | | | | | |
| Observed Incidence | - | .1 | .1 | .0 | .5 | .2 | .3 | - | 7 | |
| Adjusted Incidence | - | .3 | .3 | .0 | 1.2 | .6 | .8 | - | - | |
| Observed Prevalence | 0.0 | .1 | .2 | .0 | .5 | .2 | .9 | .0 | .0 | .1 |
| Adjusted Prevalence | - | .3 | .4 | .0 | 1.3 | .6 | 1.4 | .0 | .0 | |
| Age 40-49 | | | | | | | | | | |
| Observed Incidence | - | .7 | .0 | .8 | .0 | .7 | .0 | - | - | |
| Adjusted Incidence | - | 1.3 | .0 | 1.6 | .0 | 1.4 | .0 | .0 | _ | |
| Observed Prevalence | .9 | 1.0 | 1.1 | 2.0 | .8 | 1.4 | .8 | .7 | .9 | 1. |
| Adjusted Prevalence | - | 1.6 | 1.1 | 2.8 | .8 | 2.1 | .8 | .7 | .9 | 1. |
| Age 50-59 | | | | | | | | | | |
| Observed Incidence | - | .9 | .9 | .0 | 1.0 | .7 | .0 | - | - | |
| Adjusted Incidence | - | 1.8 | 1.8 | .0 | 2.0 | 1.4 | .0 | | - | |
| Observed Prevalence | .4 | .9 | 1.4 | .6 | 2.6 | 2.7 | 2.8 | 1.4 | .9 | |
| Adjusted Prevalence | - | 1.8 | 2.4 | .6 | 3.6 | 3.4 | 2.8 | 1.4 | .9 | |
| Age 60+ | | | | | | | | | | |
| Observed Incidence | - | 1.5 | .0 | .6 | 1.3 | .8 | .0 | - | - | |
| Adjusted Incidence | - | 3.4 | .0 | 1.2 | 3.0 | 1.9 | .0 | - | - | |
| Observed Prevalence | 2.7 | 2.6 | 2.1 | 2.4 | 3.9 | 3.3 | 3.6 | 4.0 | 2.9 | 3. |
| Adjusted Prevalence | | 4.4 | 2.1 | 3.1 | 5.5 | 4.3 | 3.6 | 4.0 | 2.9 | 3. |

APPENDIX TABLE 2 (Continued 続き)

| 1 | 2 N | 3 agasak | 4 : Eams | 5 | 6 | 7 | 8 | 9 | 10 |
|------------|------------------------------------|---|--|---|---|---|---|--|---|
| _ | N | agasak | · East | | | | NIN-S | 7 | 10 |
| - | | | rema | de | | | | | |
| - | | | | | | | | | |
| | .0 | .1 | .1 | .0 | .0 | .0 | | - | - |
| | .0 | .2 | .2 | .0 | .0 | .0 | - | = | - |
| 0. | .0 | .1 | .2 | .2 | .3 | .2 | .4 | .0 | .0 |
| 7 ° | .0 | .2 | .3 | .2 | .3 | .2 | .4 | .0 | .0 |
| | | | | | | | | | |
| - | .5 | .4 | .5 | .6 | .3 | .1 | .1 | 8 | 8 |
| - | 1.2 | 1.0 | 1.2 | 1.4 | .6 | .2 | .3 | = | H |
| .4 | .8 | .8 | .9 | 1.1 | .9 | .8 | .9 | .4 | .3 |
| - | 1.5 | 1.4 | 1.6 | 1.9 | 1.3 | .9 | 1.0 | .4 | .3 |
| | | | | | | | | | |
| - | .6 | 1.3 | .3 | 1.4 | .7 | .0 | - | 40 | - |
| - | 1.4 | 2.9 | .7 | 3.0 | 1.6 | .0 | _ | - | - |
| .5 | 1.0 | 1.6 | 1.9 | 3.6 | 3.5 | 2.4 | 2.0 | 1.8 | 1.5 |
| • | 1.8 | 3.2 | 2.3 | 5.2 | 4.4 | 2.4 | 2.0 | 1.8 | 1.5 |
| | | | | | | | | | |
| - | 1.3 | 1.0 | .0 | .6 | .6 | .2 | - | - | - |
| - | 2.7 | 2.1 | .0 | 1.2 | 1.2 | .5 | - | - | - |
| .0 | 1.3 | 2.6 | 2.4 | 2.3 | 3.7 | 4.0 | 3.9 | 3.8 | 2.7 |
| _ | 2.7 | 3.7 | 2.4 | 2.9 | 4.3 | 4.3 | 3.9 | 3.8 | 2.7 |
| | - .0 - .4 - .5 - | 0 .0 .0 0 5 - 1.2 .4 .8 - 1.5 6 - 1.4 .5 1.0 - 1.8 - 1.8 | 0 .2 .0 .0 .1 0 .2 5 .4 - 1.2 1.0 .4 .8 .8 - 1.5 1.4 6 1.3 - 1.4 2.9 .5 1.0 1.6 - 1.8 3.2 - 1.8 3.2 - 1.3 1.0 - 2.7 2.1 .0 1.3 2.6 | 0 .2 .2 .0 .0 .1 .2 0 .2 .3 5 .4 .5 - 1.2 1.0 1.2 .4 .8 .8 .9 - 1.5 1.4 1.6 6 1.3 .3 - 1.4 2.9 .7 .5 1.0 1.6 1.9 - 1.8 3.2 2.3 - 1.8 3.2 2.3 | 0 .2 .2 .0 .0 .0 .1 .2 .2 0 .2 .3 .2 5 .4 .5 .6 - 1.2 1.0 1.2 1.4 .4 .8 .8 .9 1.1 - 1.5 1.4 1.6 1.9 6 1.3 .3 1.4 - 1.4 2.9 .7 3.0 .5 1.0 1.6 1.9 3.6 - 1.8 3.2 2.3 5.2 - 1.3 1.0 .0 .6 - 2.7 2.1 .0 1.2 .0 1.3 2.6 2.4 2.3 | 0 .2 .2 .0 .0 .0 .0 .0 .0 .0 .0 .1 .2 .2 .3 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .2 .3 .2 .2 .2 .2 .2 .2 .2 .2 .2 .2 .2 .2 .2 | 0 .2 .2 .0 .0 .0 .0 .0 .0 .0 .0 .0 .0 .0 .1 .2 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .2 .3 .2 .3 .2 .2 .3 .3 .2 .3 .3 .3 .3 .3 .4 .0 .3 .3 .3 .3 .3 .3 .3 .3 .3 .3 .3 .3 .3 | 0 .2 .2 .0 .0 .0 .00 .0 .1 .2 .2 .3 .2 .40 .2 .3 .2 .3 .2 .40 .2 .3 .2 .3 .2 .4 5 .4 .5 .6 .3 .1 .1 - 1.2 1.0 1.2 1.4 .6 .2 .3 .4 .8 .8 .9 1.1 .9 .8 .9 - 1.5 1.4 1.6 1.9 1.3 .9 1.0 6 1.3 .3 1.4 .7 .0 1.4 2.9 .7 3.0 1.6 .05 1.0 1.6 1.9 3.6 3.5 2.4 2.0 - 1.8 3.2 2.3 5.2 4.4 2.4 2.0 - 1.3 1.0 .0 .6 .6 .2 2.7 2.1 .0 1.2 1.2 .50 1.3 2.6 2.4 2.3 3.7 4.0 3.9 | 0 .2 .2 .0 .0 .0 .00 .0 .1 .2 .2 .3 .2 .4 .00 .2 .3 .2 .3 .2 .4 .0 5 .4 .5 .6 .3 .1 .1 1.2 1.0 1.2 1.4 .6 .2 .34 .8 .8 .9 1.1 .9 .8 .9 .4 - 1.5 1.4 1.6 1.9 1.3 .9 1.0 .4 6 1.3 .3 1.4 .7 .0 1.4 2.9 .7 3.0 1.6 .05 1.0 1.6 1.9 3.6 3.5 2.4 2.0 1.8 - 1.8 3.2 2.3 5.2 4.4 2.4 2.0 1.8 - 1.3 1.0 .0 .6 .6 .2 2.7 2.1 .0 1.2 1.2 .50 1.3 2.6 2.4 2.3 3.7 4.0 3.9 3.8 |

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